

# College Tuition and Income Inequality\*

Zhifeng Cai

Rutgers University

Jonathan Heathcote

FRB Minneapolis and CEPR

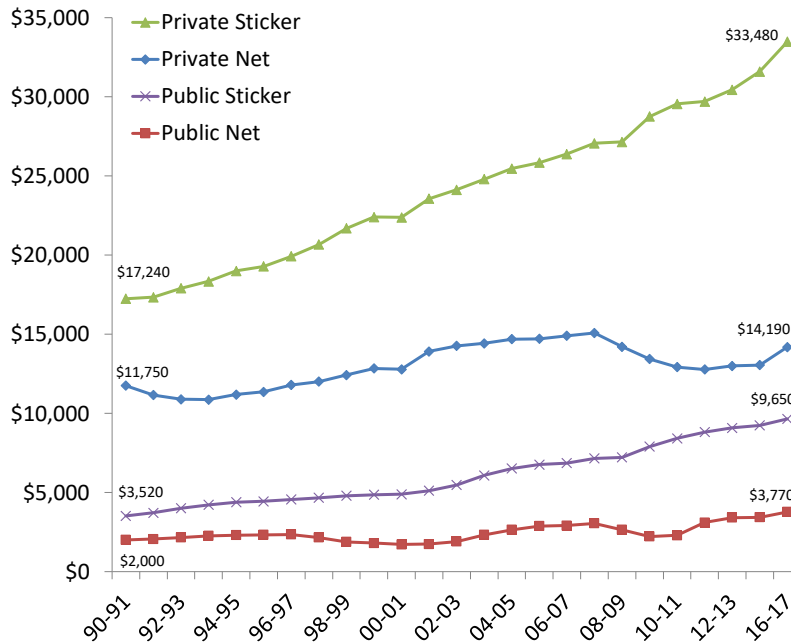
July 8 2020

## Abstract

This paper evaluates the role of rising income inequality in explaining observed growth in college tuition. We develop a competitive model of the college market, in which college quality depends on instructional expenditure and the average ability of admitted students. An innovative feature of our model is that it allows for a continuous distribution of college quality. We find that observed increases in US income inequality can explain more than half of the observed rise in average net tuition since 1990 and that rising income inequality has also depressed college attendance.

The views expressed herein are those of the authors and not necessarily those of the Federal Reserve Bank of Minneapolis or the Federal Reserve System. We thank Richard Rogerson, Grey Gordon, Sebastian Findeisen, Lance Lochner, Betsy Caucutt, John Bailey Jones, Oksana Leukhina and Todd Schoellman for helpful comments. We thank Job Boerma for research assistance. Corresponding author: Jonathan Heathcote. Mailing address: Federal Reserve Bank of Minneapolis, Research Department, 90 Hennepin Ave., Minneapolis MN 55401, USA. Tel: 612 204 6385. Email: [heathcote@minneapolisfed.org](mailto:heathcote@minneapolisfed.org)

Figure 1: College Tuition and Fees (\$2016)



Sticker and net tuition and fees for public four-year (in-state) and private nonprofit four-year colleges

## 1 Introduction

For decades, the average cost of college tuition in the United States has been rising much faster than general inflation (Figure 1), and paying for college has become a major concern for households with children. Policymakers worry that rising tuition costs may put a college education out of reach for high-ability children from low-income households. Given these concerns, it is important to understand what is driving up tuition. In this paper, we evaluate the hypothesis that rising income inequality has been a key driver of rising tuition.

This hypothesis is motivated by the fact that colleges in the United States draw their students disproportionately from relatively high-income households (see, for example, Figure 4 in [Chetty et al. 2014](#)). Rapid income growth at the top of the income distribution in recent decades has increased these households' willingness to pay for high-quality colleges. Lower-income households have experienced much weaker income growth over the same period, but this has likely not had a fully offsetting negative impact on college demand, given that few children from such households have ever attended college.

Predicting the impact of increasing college demand on college pricing requires mod-

eling the college market. The model we develop follows the existing literature in recognizing two determining factors in the quality of a college education. The first is the amount of instructional resources devoted to each student. The second is the average ability of the student body, which could be interpreted as capturing average IQ or college preparedness. Schools with higher average student ability might be more attractive to college applicants for two reasons: (i) they offer better prospects for learning from peers, and (ii) they offer social and professional connections to people who are likely to be successful post-graduation. To the extent that student ability is an important and a relatively inelastic input in producing college quality, increased demand for college will drive up equilibrium (quality-adjusted) tuition and not simply lead to an increase in the supply of high quality college spots.

Households in our model differ with respect to household income and the ability of the household child. Colleges can observe both income and ability (e.g., by observing test scores) and, in principle, can price discriminate in both dimensions. Households face tuition schedules for colleges of different quality levels and decide whether to send their child to college and, if so, to which quality of college.

On the supply side, the technology for producing college quality is a constant returns to scale function of instructional expenditure per student and average student ability. There is also a fixed cost for creating each college slot. An important feature of our model, and one that is new relative to the existing literature, is that we allow for a continuous distribution of college quality. We assume that colleges seek to provide any given value of education at the lowest possible cost or, equivalently, that they profit maximize. Colleges have no market power and thus take equilibrium tuition schedules as given. Each college chooses a quality level at which to enter and, conditional on a chosen quality level, seeks to deliver that quality as cheaply as possible by optimally balancing resource spending versus the ability composition of the student body.

As in other “club good” models, the characterization of a competitive equilibrium is complicated by the fact that club members (students) are both consumers and inputs into production, which implies a large number of market-clearing conditions. In particular, for each college quality level, the number of students demanding college spots and the ability composition of those students must be consistent with colleges’ choices about the number and composition of students to “employ” as quality-producing inputs.

Given this complication, all the existing literature in the club good tradition assumes a very small number of different college quality types. The primary theoretical contribution of our paper is to allow for a continuous distribution of quality in a competitive setting

with constant returns. This has both theoretical and practical advantages. The theoretical advantage is that we can prove an equilibrium exists. The main practical advantage is that we can compare the equilibrium model distribution of college characteristics with US data, which include thousands of different colleges. Relatedly, our continuous distribution of college quality can change smoothly when we change income inequality or other drivers of college demand.

While the model features a continuous distribution of college quality and thus a continuum of market-clearing conditions and prices, it is nonetheless quite tractable. Colleges offer lower tuition to high-ability students, internalizing that such students contribute more to college quality. We prove that this ability discount is linear in ability. There is no equilibrium price discrimination by income: any such discrimination would present an opportunity to profitably skim off high-income households. Equilibrium tuition increases with college quality, which implies a natural pattern of sorting: holding ability fixed, higher-income students match in a positive assortative fashion with higher-quality schools. Combining these insights, we show that it is possible to solve for equilibrium by iterating across the quality distribution: at each quality level (i) the density of college spots satisfies total demand, (ii) baseline tuition is such that colleges make zero profits, and (iii) the tuition discount per unit of ability equates the average ability of students wishing to attend with the average ability of students that colleges want to admit.

In the first part of the paper, we characterize equilibrium in closed form in a version of the model with no resource inputs in producing college education, two ability types, and a uniform distribution for household income. We use this closed-form example to gain intuition about what determines equilibrium college prices and the distribution by quality of college spots in a club good environment. We use it also to gain insight about how these objects vary with income inequality. The comparative statics are striking. In particular, changing income inequality has absolutely no impact on the equilibrium allocation of households across colleges of different qualities and changes only equilibrium tuition pricing.

This result motivates the second part of our paper, in which we calibrate a richer version of the model and use it to explore the role of rising income inequality and other factors in driving observed changes over time in college tuition, college attendance, and the distribution of college quality. The quantitative version of the model adds several salient features of the US college market. First, we introduce drop-out risk to reflect the fact that a large share of students who enroll in college do not graduate. We use evidence from the National Longitudinal Survey of Youth (NLSY) to calibrate how drop-out rates

vary by ability. Second, we model a range of subsidies that impact the net cost of college and thus enrollment choices. Residence-based subsidies capture tuition discounts enjoyed by in-state students enrolled at public universities. Need-based subsidies, such as Pell Grants, reduce the cost of attendance for students from low-income families. With these features, the model generates rich variation in the net cost of tuition: at any given quality level, different students face different net costs depending on ability (institutional aid), family income (need-based aid), and residence status (in-state versus out-of-state).

We set preference parameters so that the model replicates both out-of-pocket spending on college education and enrollment rates. A key model input is the joint distribution over household income and ability. We estimate a Pareto lognormal distribution for household income, using data from the Survey of Consumer Finances. We discipline the correlation between income and ability using the NLSY (where ability is proxied by Armed Forces Qualifying Test scores). The calibrated model generates a realistic enrollment pattern by family income and by student ability. It also generates distributions of sticker and net tuition across colleges that are similar to those observed empirically.

Our key quantitative experiment is to explore the implications of the change in the US household income distribution between 1989 and 2016 for the pattern of college attendance, the distribution of college quality, and the shape of equilibrium tuition schedules. Over this period, we find evidence of both a general increase in income dispersion and a significant fattening of the right tail of the distribution. These changes can account for several key features of the data. First, rising income inequality drives up average net tuition. The combined effects of rising income inequality and rising average income can account for the entire 53 percent increase in net tuition observed in the data between 1990 and 2016, with higher inequality playing the dominant role. Second, rising income inequality can account for a widening gap between average sticker tuition and average net tuition paid, reflecting increasingly large institutional discounts for desirable (high-ability) applicants (see Figure 1). Third, rising inequality depresses college enrollment, while higher average income boosts it.

We also explore the effects of changes to college subsidies and find that rising subsidies to students, in the form of more generous need-based aid and larger in-state discounts, have moderated growth in net tuition and have boosted enrollment.

**Related Literature:** There is a very large empirical literature on peer effects in education (see [Epple and Romano 2010](#) and [Sacerdote 2014](#) for excellent surveys) and a related literature that integrates peer effects into structural “club goods” models of the college market (see [Rothschild and White 1995](#) and [Epple and Romano 1998](#) for important early

contributions). In these models, and in ours, students are important inputs into the production of perceived college quality. To motivate this assumption, we now briefly discuss some empirical evidence on the importance of peers.

At a broad level, there is strong evidence that in education settings, parents and students care a lot about peers. In K–12 education, parents are willing to pay high house price premia in order to enroll their children in higher scoring and whiter school districts (e.g., [Black 1999](#); [Boustan 2012](#)). In school districts like New York and Boston there is also fierce competition to attend more academically selective magnet schools ([Abdulkadiroğlu et al. 2014](#)). At the college level, quality rankings like the ones published by US News are strongly correlated with measures of peer ability.<sup>1</sup>

Why do parents care so much about their children’s peers? The famous [Coleman \(1966\)](#) report emphasized that the educational backgrounds and aspirations of children’s peers seemed a very important predictor of student performance, while variation in resource spending across schools was not large enough to be a central factor. More recent research has tried to quantify peer effects more precisely and to uncover the mechanisms at work. [Sacerdote \(2001\)](#) found that having high GPA roommates in college tends to raise a student’s own GPA, and [Carrell et al. \(2009\)](#) find even larger effects for squadron peer groups at the US Air Force Academy. [Zimmerman \(2003\)](#) finds positive effects on grades from having high verbal SAT-scoring roommates.

It is not easy to distinguish causal effects of peers on performance from more mechanical links between individual outcomes and the average outcomes of the groups to which an individual belongs (see [Manski 1993](#) and [Angrist 2014](#)). Using a regression discontinuity approach that in principle addresses these econometric concerns, [Hoekstra \(2009\)](#) finds that attending a flagship state university boosts future earnings by 20 percent, while [Zimmerman \(2019\)](#) finds that in Chile being admitted to a highly selective college improves the chances of landing a top job. [Dale and Krueger \(2014\)](#) argue that the returns to attending a more selective college are small, but [Chetty et al. \(2020\)](#), following a similar approach, find large causal effects of colleges on future earnings, with a relatively small role for selection by unobserved student characteristics into college. Note, however, that these papers do not address to what extent better outcomes at more selective colleges reflect the presence of more desirable peers versus other factors.

Another important question is how peers matter. The typical assumption in the litera-

---

<sup>1</sup>US News explicitly assigns a ranking weight of only 10 percent to “student selectivity” (mostly SAT and ACT scores). However, the other inputs to their rankings are effectively highly correlated with these scores. [Wai et al. \(2018\)](#) report a correlation of 0.892 between college average SAT score and US News National University Rank.

ture is that average peer ability is the relevant measure (Arnott and Rowse 1987), though there is interesting recent work investigating and finding evidence that other properties of the ability distribution matter and that different types of students benefit differentially from different sorts of peers (see, e.g., Lazear 2001 and Lavy et al. 2012).

Many studies on the effects of peers have focused on relatively narrow measures of academic performance like grades and test scores. In principle, however, parents and children may care about peers for a much wider range of reasons, including friendship networks and the pool for potential spouses.<sup>2</sup> In addition, college is an important source of connections that are useful for future career development (see, e.g., Marmaros and Sacerdote 2002).<sup>3</sup> Parents and students are likely acutely aware of these sorts of peer considerations, even if they have not been the prime focus of academic research into peer effects.

On the structural modeling side, several papers build on the influential contribution of Epple and Romano (1998). In this class of models, there is a small number of colleges, which can be justified by positing large economies of scale. However, equilibrium existence problems typically arise when there are only a few college clubs, each of which is large relative to the size of the economy. These existence problems are discussed by Elickson et al. (1999) and Scotchmer (1997) and have to do with the fact that when clubs are optimally large relative to the economy, partitioning the population into an integer number of optimally sized groups is typically not possible. Because of these problems, Epple and Romano (1998) are forced to focus on approximate equilibria. They note that one could solve the existence problem by allowing for “constant costs of schooling,” which would “lead to an infinite number of schools serving infinitely refined peer groups.” They note that while such a model “is extremely interesting, it is quite complex and not yet tractable” (p. 59). This infinitely refined peer group model is the one we solve.

A second problem with assuming only a small number of competing colleges is that in such an environment, the natural model for competition is strategic oligopoly. Each college would then choose a pricing (or quality) strategy, where each strategy specifies best responses given the strategies of its competitors. Instead, most of the existing papers in

---

<sup>2</sup>For example, according to a Facebook study, at least 28 percent of college graduates marry people they meet in college: <https://www.facebook.com/notes/facebook-data-science/from-classmates-to-soulmates/10151779448773859/>. This suggests that by attending a college whose graduates earn more on average, a student can expect higher spousal earnings and thus higher household income, independent of any impact on the student’s own earnings. And if you want to marry a billionaire, you have a much better chance if you attend an elite university. Priscilla Chan met Mark Zuckerberg at Harvard, for example.

<sup>3</sup>Steve Ballmer was fortunate to be Bill Gates’ dorm hall neighbor at Harvard. Sergey Brin and Larry Page met at a new student orientation at Stanford.

the literature assume that each college takes competitors' prices (or students' willingness to pay) as given when choosing its own price. We are able to sidestep the difficult task of modeling strategic interactions between colleges: price-taking is the natural assumption in our competitive setting in which colleges are all small.

[Caucutt \(2001, 2002\)](#) takes a different approach. In her model, there is a small number of different college types but a large number of colleges of each type. Each college club is small relative to the economy as a whole, and there are no equilibrium existence problems. Households buy probabilities of attending different colleges, building on [Cole and Prescott \(1997\)](#). One limitation of her approach is that she assumes only two different income levels, which implies a very small number of different school types in equilibrium. In contrast, we assume a continuous income distribution, and we do not need to introduce lotteries (which we do not observe in practice) in order to ensure type-independent equilibrium allocations.

Two important recent papers that model the college market are [Epple et al. \(2006\)](#) and [Fu \(2014\)](#). Both papers structurally estimate rich models. Neither paper is focused on exploring the drivers of rising college tuition. In addition, there are important differences in terms of how college supply is modeled between these papers and ours. In particular, distinctive features of our model are that we allow for entry in the college market, without imposing any restrictions on the number of colleges or the qualities at which they can enter.

There is a set of papers that explores the potential drivers of rising college tuition. [Gordon and Hedlund \(2017\)](#) consider various possible factors within a variant of the model in [Epple et al. \(2017\)](#). They find that rising financial aid is the most important factor that is pushing up tuition. In their model, a single monopolistic college seeks to maximize the quality of education per student enrolled. When more public aid increases students' ability to pay, this monopolist responds by increasing spending on quality-increasing inputs. In our model, in contrast, more need-based aid induces more low-income households to enroll in college, and the market responds by expanding at the low-quality end of the distribution, thereby driving *down* average tuition.

[Jones and Yang \(2016\)](#) argue that rising college tuition reflects service sector disease: productivity in higher education is assumed to be constant, but the cost of college professors continues to rise, reflecting productivity growth and a rising college wage premium in the rest of the economy. Thus, in their model, rising income inequality plays a supply-side role in driving up the cost of college. We explore the role of rising supply-side costs and find they play a very small role in explaining tuition trends, relative to changes on



the demand side.

There is a strong positive empirical correlation between family income and college attendance and, conditional on attendance, a strong correlation between family income and proxies for college quality (see, for example, [Belley and Lochner 2007](#); [Chetty et al. 2020, 2017](#); and a recent column in the *New York Times*).<sup>4</sup> Our model predicts similar correlations. An important finding of [Belley and Lochner \(2007\)](#) is an increase in the importance of family income relative to student ability in predicting college attendance between the 1979 and 1997 waves of the NLSY. Similarly, [Chetty et al. \(2017\)](#) report in their table 2 that the share of enrollment at selective and highly selective private and public universities from households in the bottom 60 percent of the household income distribution declined from 2000 to 2011. When we feed in observed changes in the household income distribution and in college subsidies, we find that our model delivers changes in enrollment patterns similar to those reported by [Belley and Lochner \(2007\)](#) and [Chetty et al. \(2020\)](#). In particular, income becomes a more important predictor of college attendance, relative to ability, with large increases in model college enrollment from low-ability households in the top half of the income distribution.

## 2 Model

**Households:** The economy is populated by a continuum of measure 1 of households, each containing a parent and a college-age child. The baseline model is static, and within each household, the parent and child operate as a single decision-making unit. Households are heterogeneous with respect to parental income  $y$ , child “ability”  $a$ , and residence status  $r$ .

Ability is a summary statistic for any child characteristics that determine the child’s potential contribution to college quality. Ability also affects the probability of successfully graduating from university, conditional on enrolling. We assume ability is drawn from a finite discrete set of values  $A$ . Let  $\mu_a$  denote the share of the population with ability value  $a$ , with  $\sum_{a \in A} \mu_a = 1$ . Conditional on ability, household income is continuously distributed, where the cumulative distribution function  $F_a(y)$  is ability-type specific. The distribution function has a strictly positive compact support.

Colleges receive additional public subsidies for admitting in-state students, which they pass on in the form of discounted tuition. Residence status  $r \in \{i, o\}$  determines

---

<sup>4</sup>Gregor Aisch, Larry Buchanan, Amanda Cox, and Kevin Quealy, “Some Colleges Have More Students from the Top 1 Percent than the Bottom 60. Find Yours,” *TheUpshot*, *New York Times*, January 18, 2017, <https://www.nytimes.com/interactive/2017/01/18/upshot/some-colleges-have-more-students-from-the-top-1-percent-than-the-bottom-60.html>.

whether a child is eligible for these discounts:  $i$  denotes in-state, and  $o$  denotes out-of-state. The empirical counterpart to model out-of-state students will be students enrolled in private universities or in public universities as out-of-state students.<sup>5</sup> Heterogeneity in residence status is a simple and stylized way to capture heterogeneity in student access to subsidized public colleges.<sup>6</sup> The fractions of in- and out-of-state students are  $\mu_i$  and  $\mu_o = 1 - \mu_i$ . We assume that residence status is uncorrelated with parental income and child ability.

Each household chooses whether to enroll their child in college and, if so, in which quality of college. Expected utility is increasing in nondurable consumption  $c$  and in the quality of the college in which the child enrolls  $q$ . However, the utility value of college quality is realized only if the child successfully graduates, which happens with probability  $\gamma_a$ . Expected utility is given by

$$\mathbb{E}_{|a} [u(c, q)] = \log c + \varphi (\gamma_a \log(\kappa + q) + (1 - \gamma_a) \log \kappa), \quad (1)$$

where the preference parameters  $\varphi$  and  $\kappa$  are common across households. The parameter  $\kappa$  can be interpreted as the reservation utility level associated with a less-than-college education. If  $\kappa > 0$ , then utility is bounded below for children not graduating from college.

Interpreting this utility function at face value, households are willing to pay for college because it delivers a direct consumption value. Alternatively, the same utility specification can be motivated as reflecting a setting in which college is valued indirectly as an investment technology that can increase child earnings. We will further discuss this alternative interpretation in Section 6.

Households take as given tuition schedules  $t(q, y, a, r)$  that specify the out-of-pocket tuition charged by colleges offering quality  $q$ . The dependence of tuition on ability reflects the fact that different ability types are differentially attractive to colleges, and thus colleges will price discriminate. We allow tuition to depend on income in principle, but we will shortly show (Proposition 1) that any equilibrium can be supported by a tuition function that is independent of income.

---

<sup>5</sup>In a previous draft of the paper, we explicitly differentiated between private and public universities. Here we instead take the view that the key distinction in terms of tuition is between in-state students at public universities and all other students.

<sup>6</sup>There is significant variation across US states in the size of the public four-year college sector, the range of quality offered within that sector, and the size of state subsidies. State and local support for higher education in 2018 ranged from \$94 per capita in New Hampshire to \$704 per capita in Wyoming ([State Higher Education Executive Officers Association SHEF Report, 2019](#)).

Let  $\mathcal{Q}$  denote the set of college qualities available in equilibrium. One element of  $\mathcal{Q}$  is  $q = 0$ , which corresponds to the choice not to enroll in college. If the household enrolls the child in college so that  $q > 0$ , then in addition to tuition, it pays a fixed cost  $\omega$ , corresponding to the household earnings that are forgone when the child is in college instead of at work. The household can partially defray the cost of tuition through direct public financial aid,  $p(y)$ . Public financial aid depends on income, reflecting the fact that federal and grants are often means-tested. Tuition is not refunded if a student enrolls but fails to graduate. The household problem, for a household with idiosyncratic state  $(y, a, r)$  is therefore

$$\begin{aligned} & \max_{c \geq 0, q \in \mathcal{Q}} \mathbb{E}_a [u(c, q)] \\ & \text{s.t.} \\ & c + \mathbb{1}_{\{q > 0\}} [t(q, y, a, r) + \omega - p(y)] = y. \end{aligned}$$

Let  $c(y, a, r)$  and  $q(y, a, r)$  describe the decision rules that solve this problem.

**Colleges:** Colleges can enter and supply college spots at any quality level in a feasibility set  $\Omega = [0, q_{max}]$ .<sup>7</sup> The technology for producing college spots is constant returns to scale, where quality depends on the average ability of the student body and expenditure (per student)  $e$  on quality-enhancing goods and services. Each college admits a continuous mass of students. Let  $\eta_a$  denote the fraction of these students who are of ability type  $a$ . Quality (per student) at a college admitting students of average ability  $\bar{a}$  and spending  $e$  on instructional inputs is

$$q = \bar{a}^\theta e^{1-\theta}, \quad (2)$$

where  $\bar{a} = \sum_{a \in A} \eta_a a$ . The parameter  $\theta$  determines the relative importance of average student ability versus goods inputs in producing quality. Note that this production technology embeds an assumption that any mix of enrolled students with the same average ability makes an identical contribution to college quality.<sup>8</sup>

Colleges must also pay a fixed resource cost  $\phi$  per student enrolled, which captures administration and other costs that do not directly enhance quality. The fact that the technology for supplying college spots is constant returns to scale supports the existence of an equilibrium in which all colleges are small, in the sense that they enjoy no pricing

---

<sup>7</sup>The upper bound  $q_{max}$  is useful for establishing equilibrium existence. This bound can be arbitrarily large and thus will not restrict entry in equilibrium.

<sup>8</sup>Note that in the model, we have effectively defined the units of ability as measuring the marginal value of students to colleges: if one student has twice the ability of another, they are twice as valuable on the margin as a peer and twice as desirable to colleges. We will later show that an implication of this scaling choice for ability is that the equilibrium tuition function is linear in ability.

power. The logic is that each college competes against other colleges offering identical quality.<sup>9</sup>

Colleges observe the income, ability, and residence status of applicants. They seek to maximize profits or, equivalently, to provide a given market value of education at the lowest possible cost.<sup>10</sup> In addition to tuition revenue, they also receive per student public subsidies  $s(q, a, r)$  that potentially depend on the quality they deliver and on the characteristics of students they admit. In particular, colleges will receive more public support if they admit more in-state students. Note that we assume subsidies to *colleges* are independent of student income (in contrast to subsidies to *students*).

Conditional on ability, profit-maximizing colleges will strictly prefer to admit only the students who generate the most revenue. Let  $v(q, a) = \max_{y,r} t(q, y, a, r) + s(q, a, r)$  denote revenue from admitting the highest-revenue students (by income and residence status) of ability level  $a$ .

Suppose a college has decided to supply college education at quality level  $q$ . The input mix sub-problem for supplying mass one spots at quality  $q$  is

$$\begin{aligned} & \max_{\{\eta_a\} \geq 0, e \geq 0} \{ \sum_{a \in A} \eta_a v(q, a) - e - \phi \} \\ & \text{s.t.} \\ & q = \bar{a}^\theta e^{1-\theta} \text{ and } \sum_{a \in A} \eta_a = 1, \text{ where } \bar{a} = \sum_{a \in A} \eta_a a. \end{aligned}$$

Let  $\{\eta_a(q)\}_{a \in A}$  and  $e(q)$  denote the solution to this problem, and let  $\pi(q)$  denote corresponding profit per student. Given a revenue schedule  $v(q, a)$ , colleges will optimally supply zero mass of college spots at qualities  $q$  where  $\pi(q)$  is negative, will be indifferent about the mass of spots to supply if  $\pi(q) = 0$ , and will want to supply an infinite mass of spots if  $\pi(q)$  is strictly positive.

Let  $\chi(Q)$  denote the measure of college places in colleges of quality  $q \in Q \subset \Omega$ . This is a key equilibrium object. In contrast, the size distribution of colleges within any given quality level is indeterminate, given our constant returns to scale quality production function.

**Definition of Equilibrium:** An equilibrium in this model is a measure  $\chi$  and functions  $t(q, y, a, r)$ ,  $c(y, a, r)$ ,  $q(y, a, r)$ ,  $\{\eta_a(q)\}_{a \in A}$ ,  $e(q)$ , and  $\pi(q)$  that satisfy the following

<sup>9</sup>Even if a college had a monopoly at a given quality level, given a continuous quality distribution it would still face near-identical competitors and thus enjoy no pricing power.

<sup>10</sup>Some other papers in the literature assume that colleges seek to maximize college quality. In a competitive environment in which colleges take tuition schedules as given, quality maximization would imply a degenerate college quality distribution, with all colleges operating at the highest feasible quality level.

conditions:

1. Household optimization: Given  $t(q, y, a, r)$ , the household choices  $q(y, a, r)$  and  $c(y, a, r)$  solve the household's problem for all  $(y, a, r)$ .
2. College optimization: Given  $v(q, a) = \max_{y,r} \{t(q, y, a, r) + s(q, a, r)\}$ , the college input choices  $\{\eta_a(q)\}_{a \in A}$  and  $e(q)$  solve the college's problem for all  $q > 0$ , and  $\pi(q)$  is the associated profit per student.
3. Zero profits: For all  $Q \subset \Omega$ ,  $\int_Q \pi(q) d\chi(q) = 0$ , and  $\pi(q) \leq 0$  for all  $q \in Q$ .
4. Goods market clearing:

$$\sum_r \mu_r \sum_a \mu_a \int c(y, a, r) dF_a(y) + \int_0^{q_{max}} e(q) d\chi(q) + (1 - \chi(0))(\omega + \phi) = \sum_a \mu_a \int y dF_a(y) + S, \quad (3)$$

where  $S$  denotes the total value of all public subsidies

5. College market clearing: For all  $a$  and  $Q \subset \Omega$ ,

$$\mu_a \sum_r \mu_r \int \mathbb{1}_{\{q(y,a,r) \in Q\}} dF_a(y) = \int_Q \eta_a(q) d\chi(q), \quad (4)$$

where  $\mathbb{1}_{\{\cdot\}}$  is an indicator function and where for all  $y$  and  $r \in \{i, o\}$  and for all  $q^* \in Q$ ,

$$q(y, a, r) = q^* \Rightarrow (y, r) \in \arg \max \{t(q^*, y, a, r) + s(q^*, y, a, r)\}. \quad (5)$$

Condition 3 here is the zero profit condition that follows from free entry and perfect competition. It states that profits are not strictly positive at any quality level and that average profits are identically zero over any quality values at which a positive measure of college spots are supplied. Condition 4 is the goods market clearing condition. In addition to the variable cost  $e$ , each student attending college also consumes a fixed resource cost  $\omega + \phi$ . The college market clearing conditions are described in condition 5. For each ability type and for each possible set of college qualities, the number of students who wish to attend a college in that quality set must equal the corresponding number of spots supplied. And furthermore, all the students who want to attend must deliver the maximum possible revenue for colleges, conditional on ability. The fact that there are many such market-clearing conditions reflects the club good setting.

## 2.1 Equilibrium Characterization

We start by observing that any equilibrium allocation can be supported by a tuition function that does not depend on income:

**Proposition 1.** *Consider any equilibrium with an equilibrium tuition schedule  $\tilde{t}(q, y, a, r)$ . The same allocation can be supported by an alternative tuition schedule that is given by*

$$t(q, a, r) = v(q, a) - s(q, a, r), \quad (6)$$

where

$$v(q, a) = \max_{y, r} \{ \tilde{t}(q, y, a, r) + s(q, a, r) \}.$$

**Corollary.** Subsidies to in-state students pass through one-for-one into lower tuition:

$$t(q, a, o) - t(q, a, i) = s(q, a, i) - s(q, a, o) \text{ for all } q, y, a.$$

The intuition for this proposition is as follows. First, if colleges at quality  $q$  are admitting any students of ability  $a$ , they will take those delivering the highest revenue, and by definition, those students will pay  $t(q, a, r)$ . Second, if  $\tilde{t}(q, y, a, r) < t(q, a, r)$  for some idiosyncratic state  $(y, a, r)$ , then it must be the case that such households do not want to attend colleges of quality  $q$  (otherwise, condition 5 would be violated). But then the same allocation can be supported when  $\tilde{t}(q, y, a, r)$  is replaced by  $t(q, a, r)$ , because raising tuition only makes attending even less attractive. Henceforth, our measure of model tuition will refer to the function  $t(q, a, r)$  as defined above.

Note that while equilibrium sticker tuition is independent of income, low-income students still benefit from need-based financial aid via subsidies that go to students,  $p(y)$ , and thus will pay lower net tuition, all else equal.

Is this prediction of no price discrimination by income counterfactual? In practice, a portion of institutional financial aid is labeled “need-based,” which might be interpreted as income-based price discrimination. However, many colleges promise to meet the “full demonstrated financial need” of admitted students but are “need-sensitive” at the admission stage.<sup>11</sup> Among applicants who will need significant aid, these schools presumably admit only the strongest. Thus, aid that they describe as “need-based” actually has a “merit-based” component. This pattern will also arise endogenously in the equilibrium

---

<sup>11</sup>Only about 100 US colleges and universities are need blind at the admissions stage, and only half of those promise to meet full demonstrated need conditional on admission. And even within that small subset, some of this “aid” takes the form of student loans.

of our model. Because higher ability students pay lower tuition, they tend to choose higher quality colleges for any given level of household income. It follows that at any given quality level, the student body will tend to be a mix of relatively rich but low ability students who pay relatively high tuition and relatively poor but higher ability students who enjoy tuition discounts. In the context of our model, these tuition discounts truly reflect ability, but they will be systematically negatively correlated with household income, thanks to the equilibrium pattern of sorting.<sup>12</sup>

**Proposition 2.** *Any equilibrium can be supported by a revenue schedule that is linear in ability, that is, one that takes the form*

$$v(q, a) = b(q) - d(q)(a - a_{\min}),$$

where  $a_{\min}$  is the lowest value in the set  $A$ . The quality-dependent constant  $b(q)$  defines revenue from the lowest ability type, while  $d(q) > 0$  denotes the revenue discount per unit of ability.

Proposition 2 greatly simplifies equilibrium characterization. A linear revenue function ensures that the revenue from admitting any set of students is linear in the average ability of the students in that set. Thus, we can rewrite the college problem as

$$\begin{aligned} \max_{\bar{a}, e} \{ & b(q) - d(q)(\bar{a} - a_{\min}) - e - \phi \} \\ & \text{s.t.} \\ & q = \bar{a}^\theta e^{1-\theta} \\ & a_{\min} \leq \bar{a} \leq a_{\max}, \end{aligned} \tag{7}$$

where  $\bar{a}$  is the average ability of students admitted.<sup>13</sup> The first-order condition is

$$\frac{d(q)}{1} = \frac{\theta \frac{q}{\bar{a}}}{(1-\theta) \frac{q}{e}}. \tag{8}$$

The left-hand side is the ratio of the price of a marginal increase in average ability relative to the price of a marginal increase in  $e$ . The right-hand side is the corresponding ratio of marginal products. If the equilibrium revenue schedule declines steeply with

---

<sup>12</sup>See [Chetty et al. \(2020\)](#) Online Appendix Table XII for evidence of the same sorting pattern by income and ability within “Ivy Plus” colleges.

<sup>13</sup>Note that  $\bar{a}$  is a continuous choice variable, even though  $a$  takes a discrete set of values because each college enrolls a continuum of students and can vary average student ability smoothly by varying the enrollment shares  $\eta_a$ .

ability ( $d(q)$  is large), then colleges will choose a high ratio of instructional inputs relative to average student ability.<sup>14</sup>

We now turn to households' choices about college attendance and quality. Absent need-based financial aid, the optimal choice for quality is increasing in household income, holding fixed ability. This is because college quality is a normal good and because equilibrium tuition, as just discussed, is independent of income. This property simplifies equilibrium computation because it means that moving up the college quality distribution, college spots for each ability type will be filled in a strictly ordered fashion by income, with the highest quality colleges taking students from the top of the income distribution.<sup>15</sup>

## 2.2 Existence and Pareto Efficiency

We now discuss equilibrium existence and efficiency. For ease of exposition, in this section we focus on a version of the model that abstracts from government intervention.

Existence has been a long-standing issue in college models with peer group effects (Epple and Romano, 1998, 2010). In our environment, with a constant returns technology and a continuum of agents, we are able to prove a formal existence result, one that does not rely on introducing lotteries.<sup>16</sup>

**Proposition 3.** *A competitive equilibrium exists.*

We next prove that the competitive equilibrium is Pareto efficient. A crucial assumption underlying this result is that clubs are competitive price takers. This implies that the peer group effect is a "local" externality within a college and is correctly priced in the

---

<sup>14</sup>One can further show, with the zero-profit condition, that the revenue function  $v(\cdot)$  must satisfy effective marginal cost (EMC) pricing as in Epple and Romano (2008). That is, the revenue function must consist of three components: a fixed cost component; a variable expenditure component; and a component reflecting the student's peer group value:

$$v(q, a) = \underbrace{\phi}_{\text{Fixed Cost}} + \underbrace{e(q)}_{\text{Variable Cost}} + \underbrace{\frac{\theta}{(1-\theta)} \frac{e(q)}{\bar{a}(q)} (a - \bar{a}(q))}_{\text{Peer Value}},$$

where  $e(q)$  and  $\bar{a}(q)$  are solutions to the colleges' problem. This function is unique to admitted students.

<sup>15</sup>Need-based financial aid potentially breaks this result, depending on the schedule for  $p(y)$ .

<sup>16</sup>The proof of existence consists of two steps. In the first step, we establish that a competitive equilibrium exists when there are finitely many different types of college. This part of the proof draws heavily on Ellickson et al. (1999). The second step extends the existence result to a continuum of club types. To do so, we adopt the methodology from Caucutt (2001) and construct a sequence of approximate economies with finitely many college types, and show that the limit of that sequence is the equilibrium for the economy with a continuum of college types.



competitive market: higher-ability students are charged lower tuition. Thus, as argued by [Ellickson et al. \(1999\)](#), a club goods economy is conceptually no different from a non-club goods economy, in the sense that (type-specific) club memberships can be treated as ordinary goods traded in a competitive market.

**Proposition 4.** *A competitive equilibrium, absent government subsidies, is Pareto efficient.*

The proof here closely follows the standard proof of the First Welfare Theorem. The result hinges critically on the assumption that households care about quality directly and derive utility only from their own child’s college quality, so college is a pure consumption good. High-income households will spend a lot on tuition at expensive schools – even if their children are low ability – because their marginal utility from non-college consumption is low. Thus, they have a high willingness to pay for the experience and prestige associated with high quality colleges, or for gaining access to a more attractive pool of potential spouses and professional connections. Similarly, if poor households choose not to send their children to college, that will simply reflect the fact that those households prefer to spend their limited income on consumption goods.

In Section 6, we discuss an alternative interpretation of the model, in which the value of college has an investment component. In this case, allocations are *not* efficient.

### 3 A Closed-Form Example

Before calibrating the model described above, we first show that in one special case, equilibrium allocations can be characterized in closed form. This example is useful because it clearly illustrates that the club good nature of the college market has important implications for college pricing and for the effects of changes in income inequality on the allocation of students to colleges and the tuition they pay. It also provides a proof of existence by construction in this special case.

The special case is one in which  $\theta = 1$ , so average student ability is the only determinant of college quality. In addition,  $\omega = \phi = 0$ , so there are no fixed resource costs associated with attending college. The preference parameter  $\varphi$  is equal to one. There are no public subsidies of any form. There are two ability types, denoted by  $a_l$  and  $a_h$ , and each half the population is of one type. The income distribution is independent of ability and uniform:  $y \sim \mathcal{U}(\mu_y - \frac{\Delta_y}{2}, \mu_y + \frac{\Delta_y}{2})$ , where  $\mu_y$  denotes average income and  $\Delta_y$  defines income dispersion. Let  $\Delta_a = a_h - a_l$  denote ability dispersion and  $\mu_a = \frac{a_h + a_l}{2}$  denote average ability.

Note that given  $\theta = 1$ , the production function implies that  $q = \bar{a}$ . Note also that the support of possible college qualities is  $[a_l, a_h]$ .

**Proposition 5.** *Under the parameterization described above, the model has a competitive equilibrium in which the measure of college spots by quality and tuition schedules are described, respectively, by*

$$\begin{aligned}\chi(Q) &= \frac{2}{\Delta_a} \frac{2}{(4 + \pi)} \int_Q \left[ \left( \frac{a_h - q}{\Delta_a} \right)^2 + \left( \frac{q - a_l}{\Delta_a} \right)^2 \right]^{-2} dq \quad \forall Q \subset (a_l, a_h), \\ \chi(a_h) &= \chi(a_l) = \frac{2}{4 + \pi}, \\ t(q, a_i) &= \mu_y \frac{1}{\kappa + q} \left[ 1 - \frac{2}{4 + \pi} \frac{\Delta_y}{\mu_y} \arctan \left( 2 \frac{(\mu_a - q)}{\Delta_a} \right) \right] (q - a_i),\end{aligned}\tag{9}$$

where  $\pi$  is the mathematical constant and  $\arctan$  is the inverse tangent function.

The equilibrium tuition function satisfies what [Epple and Romano \(1998\)](#) term “effective marginal cost pricing”:

$$t(q, a_i) = \sum_j \eta_j \frac{\partial t(q, a_j)}{\partial q} (q - a_i).$$

This optimality condition equates the tuition revenue from admitting a student of ability  $a_i$  to the marginal cost of admitting such a student via the peer group externality. In particular, admitting a lower than current average ability student lowers college quality proportionally to  $q - a_i$  and thus lowers the tuition the college can charge to all students.<sup>17</sup>

<sup>17</sup>See [Epple and Romano \(1998\)](#) (page 40) for more discussion. To see that this condition is satisfied, note that the equilibrium tuition function (eq. 9) can be written as  $t(q, a_i) = G(q)(q - a_i)$  where

$$G(q) = \mu_y \frac{1}{\kappa + q} \left[ 1 - \frac{2}{4 + \pi} \frac{\Delta_y}{\mu_y} \arctan \left( 2 \frac{(\mu_a - q)}{\Delta_a} \right) \right].$$

We now establish that  $\sum_j \eta_j \frac{\partial t(q, a_j)}{\partial q}$  is equal to  $G(q)$ .

$$\begin{aligned}\sum_i \eta_i \frac{\partial t(q, a_i)}{\partial q} &= \sum_i \eta_i (G'(q)(q - a_i) + G(q)) \\ &= \frac{q - a_l}{\Delta_a} [G(q) + (q - a_h) G'(q)] + \frac{a_h - q}{\Delta_a} [G(q) + (q - a_l) G'(q)] \\ &= \frac{a_h - a_l}{\Delta_a} G(q) = G(q),\end{aligned}$$

where the first equality substitutes in  $t(q, a_i) = G(q)(q - a_i)$  and differentiates with respect to  $q$ ; the second

The equilibrium allocations described in equation (9) have several interesting properties. First, the quality distribution of college spots is independent of the income distribution parameters  $\mu_y$  and  $\Delta_y$  and is also independent of the preference parameter  $\kappa$ . Given perfect assortative matching from income to quality, this result implies that the equilibrium college quality choice for a household depends only on the household *rank* in the income distribution and their child's ability. The quality distribution of college spots has two mass points at the lowest and highest possible quality colleges,  $q = a_l$  and  $q = a_h$ . In between these values, the distribution is continuous, symmetric, and single peaked. The lowest quality schools are filled with low-ability students drawn from the bottom of the income distribution, while the highest quality schools are filled with high-ability students drawn from the top. Outside these income ranges, students attend mixed-ability schools.

In contrast, the income distribution parameters  $\mu_y$  and  $\Delta_y$  do appear in the equilibrium tuition functions. Tuition is an increasing but nonlinear function of quality.<sup>18</sup> The tuition ability discount  $d(q) = \frac{t(q, a_l) - t(q, a_h)}{\Delta_a}$  is increasing (decreasing) in income dispersion  $\Delta_y$  for  $q \geq \mu_a$  ( $q \leq \mu_a$ ).

Figures 2 and 3 plot  $\chi(q)$  and  $t(q, a_i)$  for an example in which  $a_l = 0$ ,  $a_h = 1$ ,  $\mu_y = 1$ , and  $\kappa = 5$ . We consider two different values for income inequality: low inequality, with  $\Delta_y = 0.1$ ; and high inequality, with  $\Delta_y = 1.9$ . The allocation of students to colleges, depicted in Figure 2, is exactly the same in both cases. When income dispersion is increased, Figure 3 indicates that low-ability students pay higher tuition at high quality colleges, while high-ability students at the same colleges enjoy larger tuition subsidies (recall that average tuition, weighted by ability, is always equal to zero, thanks to free entry). One could interpret this pattern as an increase in sticker tuition, coupled with greater merit aid for high-ability students, such that the gap between sticker tuition and net tuition widens.

Why is the distribution for college quality  $\chi$  completely insensitive to income inequality? A partial intuition is that the support of the college quality distribution cannot expand in our environment: the bounds are always  $q = a_l$  and  $q = a_h$ . Clearly, in order for households at the extremes of the income distribution to choose feasible values for quality, tuition schedules must move when income inequality is increased. The fact that the richest low-ability households are now richer increases the relative demand from low-ability households for high quality colleges. In equilibrium, a rise in low-ability tuition at high quality colleges induces these richer households to leave their quality choices unchanged.

---

equality utilizes the admission rules for high and low ability students  $\eta_h = \frac{q - a_l}{\Delta_a}$  and  $\eta_l = \frac{a_h - q}{\Delta_a}$ ; the third equality cancels out terms involving  $G'(q)$ ; and the fourth equality follows from the definition  $\Delta_a = a_h - a_l$ .

<sup>18</sup>It is easy to check that  $t(q, a_l) \geq 0$  and  $t(q, a_h) \leq 0$  for all  $q \in [a_l, a_h]$ .

Figure 2: College Distribution

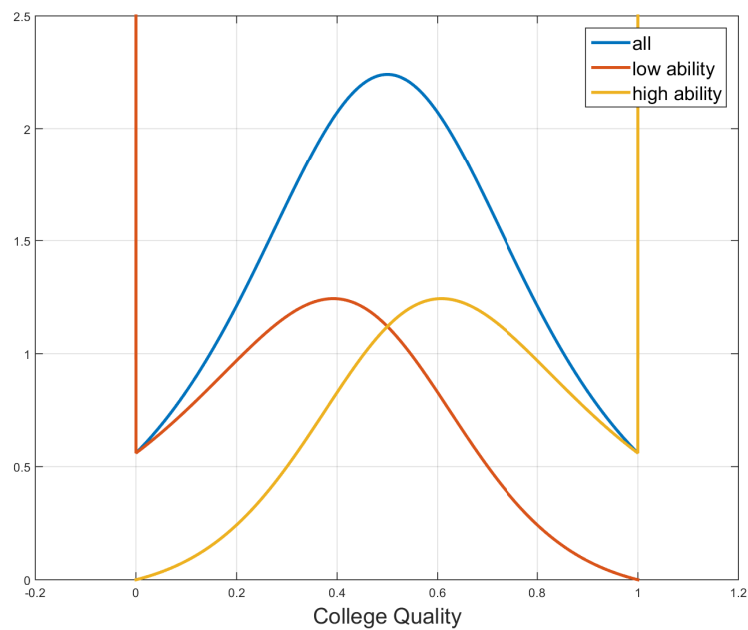
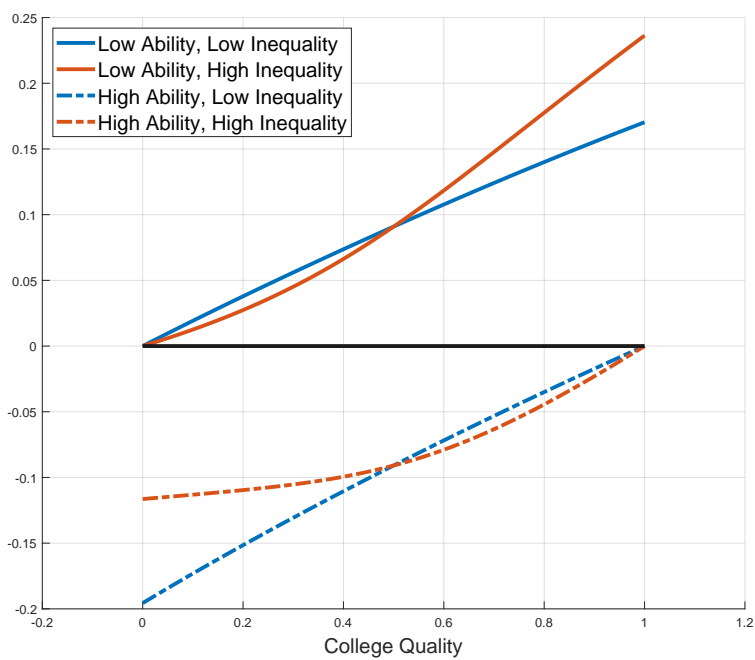


Figure 3: College Tuition Schedules



Similarly, rising income inequality leaves the poorest high-ability students poorer. A rise in high-ability tuition at low quality colleges induces these households not to downgrade college quality.

Note that the implications of widening inequality in this club good setting are starkly opposite to those in a conventional non-club goods model. In a conventional model, increasing income dispersion would lead the rich to increase consumption of college quality, while the poor would reduce quality, and the price per unit of quality would remain unchanged. In the club good model, in contrast, all the effects of widening income inequality show up in changes in equilibrium tuition, with no impact on the allocation of college quality. Thus, this simple example offers a powerful motivation for exploring the potential role of widening income inequality on college tuition in a richer calibrated model, to which we now turn.

## 4 Quantitative Application

Our baseline calibration is for 2016.<sup>19</sup> We assume that the distribution for household income (conditional on child ability) is Pareto lognormal, a parametric functional form that closely approximates the actual distribution of income in the United States (see [Heathcote and Tsujiyama 2019](#)). Thus, log household income is given by  $\log y = x_1 + x_2$ , where  $x_1$  and  $x_2$  are independent random variables,  $x_1$  is normally distributed with mean  $\mu_y$  and variance  $\sigma^2$ , and  $x_2$  is exponentially distributed with exponential (Pareto index) parameter  $\alpha$ . This distribution transitions smoothly from an approximately lognormal distribution over most of the income distribution toward a Pareto distribution in the right tail. We estimate the parameters  $\mu_y$ ,  $\sigma^2$ , and  $\alpha$  using microdata on log total household income from the 2016 Survey of Consumer Finances. One important strength of this survey is that households at the top of the income distribution are not underrepresented, which is important for being able to estimate the Pareto parameter  $\alpha$ . Because we are interested in income for households that are making decisions about college, we restrict our sample to households between ages 40 and 59. The maximum likelihood estimates for  $\sigma^2$  and  $\alpha$  are 0.55 and 1.67, implying a variance of log income equal to  $0.55 + 1.67^{-2} = 0.91$ .<sup>20</sup>

As in the closed-form example, our baseline calibration assumes two ability types.<sup>21</sup>

---

<sup>19</sup>See the Data Appendix for more details on the construction of the statistics used in calibration and model-data comparisons.

<sup>20</sup>These estimates are similar to those from other studies (e.g., [Piketty and Saez, 2003](#)). Measured income dispersion may actually understate true income dispersion, thanks to non-classical measurement error in survey data ([Gottschalk and Huynh, 2010](#)).

<sup>21</sup>We develop a specialized computational algorithm for the two-ability type case, which is much faster

We assume that the income distributions conditional on ability are both Pareto lognormal, with the same estimated values for  $\sigma^2$  and  $\alpha$ . To allow for correlation between household income and child ability, we index the level parameter by ability, setting  $\mu_{a_h} = \mu_y + \delta$  and  $\mu_{a_l} = \mu_y - \delta$ . To estimate  $\delta$ , we turn to the 1997 NLSY and use AFQT scores as a proxy for child ability. We rank children by these scores and set  $\delta$  to replicate the ratio of average family income for households with children in the top versus the bottom half of the AFQT score distribution, which is 1.59.<sup>22</sup>

Hendricks et al. (2018) report graduation rates (i.e., one minus dropout rates) by AFQT score, conditional on enrollment in four year colleges, again using the 1997 NLSY. These rates are 78 percent for students in the top half of the AFQT distribution and 52 percent for students in the bottom half. Thus, we set the graduation rate probabilities to  $\gamma_{a_h} = 0.78$  and  $\gamma_{a_l} = 0.52$ .

We set the reservation utility parameter,  $\kappa$ , to target a model college enrollment rate of 50.7 percent. Given the graduation probabilities described above, this is the enrollment rate consistent with the observed mix of above and below median ability students graduating from four year colleges, as well as with an aggregate graduation rate of 36.1 percent, which is the share of individuals aged 25–29 who report at least a bachelor’s degree in the CPS.

The remaining model parameters are set to replicate various moments involving tuition, financial aid, subsidies and costs for the universe of nonprofit four-year private and public colleges.<sup>23</sup> Data on tuition, aid, and costs are from the College Board and the National Center for Education Statistics (NCES). Many statistics are reported separately for public and private four-year colleges. Recall that we do not separately model the two sectors. Rather we focus on the distinction between public students who are in-state and all other students, taking the view that private students and public out-of-state students are effectively active in the same market. We assume that universities receive additional government subsidies in proportion to the number of in-state students that they admit and therefore offer those students discounted tuition. In 2017, 70 percent of students in four-year colleges were in public universities, and of those, 78 percent were in-state students. We therefore set  $\mu_i$ , the share of in-state students in the population, to target a

---

than our general algorithm that handles more ability types. In Appendix 8.2.2, we show that introducing more ability types has a negligible impact on the distribution of college enrollment and tuition.

<sup>22</sup>We set  $\mu_y$  to target a value for average income,  $\bar{y} = 0.5 \left[ \exp\left(\mu_y + \delta + \frac{\sigma^2}{2}\right) \frac{1}{\alpha} + \exp\left(\mu_y - \delta + \frac{\sigma^2}{2}\right) \frac{1}{\alpha} \right]$ .

<sup>23</sup>Around one-tenth of individuals with a bachelor’s degree are graduates from for-profit colleges, which could be conceptualized as low-quality private schools in the context of our model. However, because of data limitations, we exclude the for-profit sector when computing the statistics targeted in our calibration.

share of in-state students in the model of  $0.7 \times 0.78 = 54.6$  percent. The implied  $\mu_i$ , is 0.529. This is slightly lower than the target enrollment share because subsidized in-state students are more likely to choose to enroll.

The preference parameter  $\varphi$  is set to replicate average college tuition paid. We focus on net tuition, defined as sticker tuition minus government and institutional aid. This reflects the average amount students pay out of pocket and is therefore a good gauge of the strength of the preference for college. Average net tuition for the 2016–17 academic year was \$3,770 for public in-state students, \$14,190 for private university students, and \$19,050 for public out-of-state students.<sup>24</sup> Given the respective enrollment shares for the three types of students, average net tuition was \$9,250.

We next translate this dollar value into an estimate of net tuition at four year colleges as a share of total consumption spending, which is an object we can readily compute in both model and data. In 2016, there were 8.34 million students enrolled in four-year schools out of a total US population of 323.4 million, implying an aggregate attendance rate of 2.58 percent.<sup>25</sup> In 2016, private consumption per capita was \$39,417. Thus, aggregate net tuition spending on four-year colleges was  $2.58 \times \frac{9,249}{39,417} = 0.61$  percent of total consumption.<sup>26</sup> We set  $\varphi$  so that model average net tuition times the target enrollment rate (0.507) is equal to 0.61 percent of average model income.<sup>27</sup>

Forgone earnings  $\omega$  is an important additional component of the cost of attending college. We set  $\omega$  equal to 20 weeks of median weekly earnings for full-time workers aged 16 to 24, which was \$10,020 in 2016 (Current Population Survey).<sup>28</sup> We exclude room and board from our measure of college prices, on the grounds that similar living

---

<sup>24</sup>We observe average sticker tuition for public out-of-state students but not average net tuition. We estimate average net tuition for this group by assuming that the sticker versus net tuition differential for out-of-state students is the same as for public in-state students.

<sup>25</sup>Given the observed 50.7 percent enrollment rate, and assuming the U.S. economy is in steady state, this translates into each college enrollee spending  $0.0258/0.507 = 5.1$  percent of their lifetime in college, or 4.1 years, assuming an 80 year lifetime.

<sup>26</sup>The share of college tuition and fees in the Consumer Price Index is larger at 1.8 percent. However, the CPI college category includes two-year schools and graduate and professional schools. Also, the CPI tuition measures are closer to capturing sticker price than actual price paid (e.g., [Schwartz and Scafidi 2004](#)).

<sup>27</sup>Average model income is the value for  $\bar{y}$  that solves  $\frac{0.507 \times 9,250}{\bar{y}} = 0.0258 \times \frac{9,250}{39,417}$ , which implies  $\bar{y} = \$774,590$ . The reason this is a large number is that in our static model, the ratio of tuition to income should be thought of as average lifetime tuition to average lifetime income: a student with average household income attending college and paying average net tuition will spend  $9,250/774,590 = 1.2\%$  of household income on tuition.

<sup>28</sup>We chose 20 weeks, thinking that students can work full-time when school is out (typically 16 weeks a year), and part-time during the school year. The [College Board \(2016, p. 8\)](#) suggests using living expenses as a proxy for the cost of forgone earnings. The average cost of room and board at 4-year colleges in 2016–17 was \$10,875, which is similar to our earnings-based estimate.

expenses apply irrespective of college attendance.<sup>29</sup>

We next describe the model for student financial aid,  $p(y)$ . We assume that this aid has a general component  $p_0$  and a need-based component  $p_1$  that is available to students with family income below a threshold  $y^*$  :

$$p(y) = p_0 + \mathbb{1}_{\{y \leq y^*\}} p_1.$$

The main form of federal need-based aid is Pell Grants. In 2016–17, 32 percent of undergraduate students received a Pell Grant, and the average grant size was \$3,800. A large share of state grant aid is also need-based, translating into \$982 per student. We assume the students receiving need-based state grants are the same as those receiving Pell Grants, implying  $p_1 = 3,800 + 982/0.32 = \$6,870$ . We then set the threshold  $y^*$  such that in equilibrium, 32.0 percent of enrolled students come from families with income below  $y^*$ , which requires  $y^*$  equal to 71.4% of mean model income.<sup>30</sup>

We pin down the general component of the subsidy  $p_0$  residually, exploiting the accounting identity that the sum of average government aid plus institutional aid is equal to the difference between average sticker tuition and average net tuition. Average sticker tuition in 2016–17 was \$19,152 (\$9,650 for public in-state students, \$33,480 for private students, and \$24,930 for public out-of-state students). Our estimate for average institutional aid per student is \$5,808 (see below). Thus,  $p_0 = \$19,152 - \$9,250 - \$5,808 - 0.32 \times \$6,870 = \$1,896$ .

We turn next to college subsidies. Our model for the form of these subsidies is

$$\begin{aligned} s(q, a, o) &= \bar{s} \\ s(q, a, i) &= \bar{s} + \max \{ (1 - \lambda)t(q, a, o), 0 \}. \end{aligned}$$

Thus, colleges receive a fixed per student subsidy  $\bar{s}$  for every student they enroll. In addition to various forms of government support, this captures any other source of non-tuition income, such as endowment income and private gifts. Colleges that admit in-state students receive an additional per student subsidy, which is a fraction  $(1 - \lambda)$  of the equilibrium tuition that a similar out-of-state student would pay at a school of the same

---

<sup>29</sup>Increased top tail income inequality will translate into greater demand for high quality accommodation and food and thus can potentially rationalize observed growth over time in the cost of room and board.

<sup>30</sup>Our model for need-based aid is simpler than the actual Pell grant program. The size of Pell grant a student can expect depends on the student's Expected Family Contribution (which depends on family income) and also on the cost of college attendance. However, there is an upper bound on the maximum possible grant, which was \$5,920 in 2016.



quality. Note that given eq. 6, this model implies that in equilibrium, in-state students pay a constant fraction  $\lambda$  of what identical out-of-state students pay:  $t(q, a, i)/t(q, a, o) = \lambda$ .<sup>31</sup>

Average sticker tuition for public out-of-state students in 2016–17 was \$24,930, while the corresponding number for in-state students was \$9,650, suggesting an average extra subsidy per in-state student of \$15,280 (or \$8,343 per student overall). We set  $\lambda$  to replicate this average per student subsidy across all in-state students, which implies  $\lambda = 0.49$ .

Next, we turn to the ratio between ability levels,  $a_h/a_l$ , and the share parameter,  $\theta$ , which defines the relative importance of average student ability versus instructional spending in determining college quality. These parameters jointly determine the importance of the club good feature of the model. Increasing  $a_h/a_l$  or increasing  $\theta$  both make high ability students more attractive for colleges, which translates into more institutional aid being used to recruit such students. Average institutional aid per student in 2015–16 was \$2,274 at public four-year institutions and \$14,055 at private institutions, for an overall average of \$5,808. We normalize  $a_h = 1$  and set  $a_l$  so that the model replicates this value, which implies  $a_l = 0.375$ .<sup>32</sup> Note that in the model, all low ability students pay full sticker tuition, while all institutional aid goes to high ability students. Because  $a_h/a_l$  and  $\theta$  are not sharply separately identified, following much of the existing literature we simply set  $\theta = 0.5$ , implying that peer effects and expenditure are equally important in delivering quality (see, e.g., [Caucutt 2002](#) and [Epple et al. 2006](#)).<sup>33</sup> We will also experiment with  $\theta = 0.25$  and  $\theta = 0.75$  in Section 5.4.

The only remaining parameters are the fixed cost for admitting students  $\phi$  and the fixed per student subsidy  $\bar{s}$ . From the college problem, it is clear that colleges care only about the net fixed cost of enrolling students,  $\phi - \bar{s}$ . In equilibrium, colleges make zero profits. We can therefore estimate the net fixed cost using the following equation equating

---

<sup>31</sup>From the College Scorecard data, we find that the empirical ratio of sticker tuition for in-state versus out-of-state students does not systematically vary with the *level* of tuition across public universities, consistent with this subsidy model.

<sup>32</sup>The units of ability and income in the model are both arbitrary, and thus we are free to choose the units of each. Given parameter values and a particular choice of units, one can characterize a competitive equilibrium. It is straightforward to see how the same equilibrium must be rescaled under alternative choices for units. For example, if we multiply the units for the output / consumption good by a factor  $m$ , then the same equilibrium is preserved as long as the preference parameter  $\kappa$  is scaled by a factor  $m^{1-\theta}$ . Similarly, if we multiply the units for ability by a factor  $n$ , the same equilibrium is preserved as long as  $\kappa$  is scaled by a factor  $n^\theta$ .

<sup>33</sup>The standard deviation of test scores on standardized tests cannot be used to identify the ratio  $a_h/a_l$  since test score dispersion is a choice of the test designer, rather than an informative empirical moment that can be targeted.

Table 1: Calibration

Income distribution		Aid	
$\mu_y$	13.77	Residence-based	
$\delta$	0.232	$\mu_i, \mu_o$	0.529, 0.471
$\sigma^2$	0.548	$\lambda$	0.490
$\alpha$	1.670	Need-based	
Preferences		$y^* / \bar{y}$	0.714
$\varphi$	0.049	$p_1$	\$6,870
$\kappa$	36.89	General	
Technology		$p_0$	\$1,896
$\theta$	0.50	$\phi - \bar{s}$	\$4,610
$a_h, a_l$	1, 0.375		
$\gamma_{a_h}, \gamma_{a_l}$	0.781, 0.520		
$\omega$	\$10,020		

college sector income to expenses:

$$\underbrace{E[t(q, a, r)]}_{\$19,152-\$5,808} + \underbrace{E[s(q, a, r) - \bar{s}]}_{\$8,343} = \underbrace{E[e(q)]}_{\$17,077} + \underbrace{(\phi - \bar{s})}_{\$4,610}.$$

The first term on the left-hand side is average tuition income (average sticker tuition minus average institutional aid). We identify variable resource spending  $e$  with the sum of the NCES expenditure components “instruction” and “student services.” In 2016, instruction and student services spending per student averaged \$22,749 at private four-year schools and \$14,646 at public schools, for an overall average of \$17,077 per student. From the equation above, the implied net fixed cost is \$4,610. The calibration is summarized in Table 1.

The model is computed on a discretized grid of college quality. The upper bound  $q_{max}$  is set large enough that it does not restrict entry of colleges at the top. We use 400 grid points for quality, and we have checked that making the grid finer does not affect equilibrium prices or allocations. We set the maximum level of income to be 20 times the average, which corresponds to the 99.8<sup>th</sup> percentile of the SCF income distribution. We develop an iterative algorithm that exploits the assortative pattern of sorting by income to college quality, conditional on ability, residence status, and need-based aid eligibility. A detailed description of the algorithm is in the Appendix.

## 4.1 Model Predictions

To assess whether ours is a reasonable model of the college sector, we examine what it implies for some key non-targeted moments. In particular, does the model successfully replicate who enrolls in college, and how much tuition they pay?

In terms of enrollment, we focus on the income and ability composition of students enrolling in college (we match the in-state versus out-of-state mix by construction). [Chetty et al. \(2020\)](#) report average parental family income at the college level, constructing these estimates from federal income tax records.<sup>34</sup> Average parental income for children in four-year universities is 56.0 percent larger than the average for all children. The corresponding statistic for the model is very similar at 56.7 percent, implying that the model delivers realistic sorting into college by family income. Using the NLSY 97, we estimate that 74.9 percent of above median ability children enroll in college, compared with 26.5 percent of below median ability children. The corresponding enrollment rates in the model are 80.2 percent and 21.2 percent, indicating that the model does a good job replicating the pattern of college enrollment by ability.<sup>35</sup>

Next, we compare the model's implications for the cross-sectional distribution of both sticker and net tuition. Our calibration procedure ensures that the model matches the empirical values for average sticker and net tuition, but the pattern of dispersion in tuition across schools of different qualities is endogenous and untargeted. Data on tuition at the college level are available from the College Scorecard, which is a tool provided by the US Department of Education to facilitate comparison shopping across colleges.<sup>36</sup>

Figure 4 plots the distributions of net and sticker tuition, model against data. Overall, the distributions look broadly similar. One discrepancy is at the very top of the distribution, where published tuition and fees at elite private universities top out at around \$55,000 in the data, while the model predicts a right tail of even more expensive colleges.<sup>37</sup> Table 2 indicates that the coefficients of variation for net and sticker tuition across

---

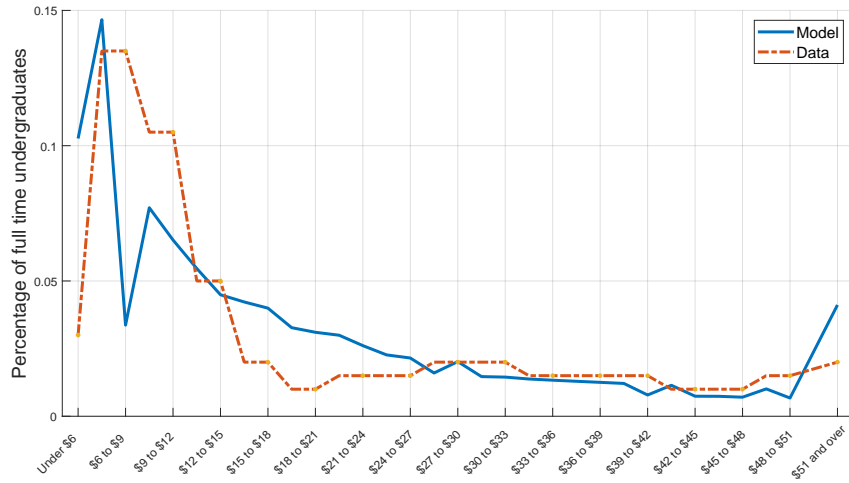
<sup>34</sup>Parental income is averaged over 1996–2000, a five-year period in which the child (and potential college attendee) is aged 15–19.

<sup>35</sup>Recall that our calibration targets the aggregate enrollment rate, and that graduation rates are higher for high ability students. Because the model slightly over-predicts (under-predicts) high ability (low ability) enrollment, it delivers a slightly higher graduation rate than the data: 36.9 percent versus 36.1 percent.

<sup>36</sup>The College Scorecard reports sticker tuition and fees and the average net price of attendance for undergraduates receiving Title IV aid, which is the full cost of attendance less federal, state, and institutional grant aid. This measure includes living expenses. To construct a net tuition measure that excludes living expenses, we estimate living expenses at the college level as the difference between the average annual full cost of attendance and tuition and fees.

<sup>37</sup>Note, however, that wealthy families at elite private colleges are informally expected to make voluntary donations on top of paying published tuition. Thus, published tuition arguably understates the true

(a) Sticker Tuition Distribution (2016, \$1,000)



(b) Net Tuition Distribution (2016, \$1,000)

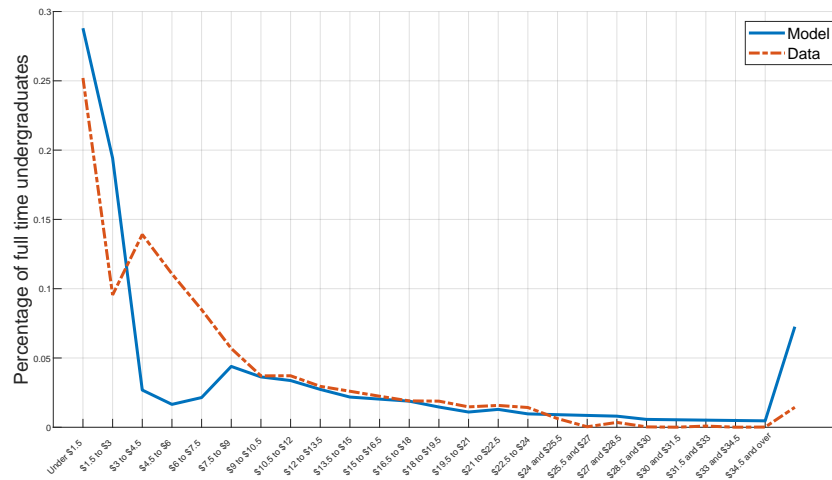


Figure 4: Sticker and Net Tuition Distributions

all four-year colleges in the United States are 0.99 and 0.77. The analogous statistics from the model are 1.31 and 0.80.

The model also generates a fairly realistic joint distribution across colleges, in terms of other observable college characteristics: average household income of students attend-

---

cost of attendance for these families. Applications from the wealthiest families are typically known as “development cases” and are handled separately from regular admissions, with enrollment offers more or less explicitly tied to actual or expected donations (see Golden 2006). For such students, sticker price may vastly understate the true cost of admission.

Table 2: Non-Targeted Moments Comparison

	Data	Model
<i>Enrollment Patterns</i>		
Family income enrolled / Mean	1.560	1.567
Share of high ability enrolled	0.749	0.802
Share of low ability enrolled	0.265	0.212
Graduation Rate	0.361	0.369
<i>College-level Moments</i>		
Standard Deviation / Mean		
Net tuition	0.99	1.31
Sticker tuition	0.77	0.80
Avg. family income	0.51	0.92
Fraction of high ability	0.26	0.10
Correlation		
Sticker tuition vs. Net tuition	0.83	0.98
Net tuition vs. Family income	0.60	0.97
Net tuition vs. Fraction of high ability	0.22	0.71
Family income vs. Fraction of high ability	0.59	0.77

ing, and the share of high-ability students enrolled.<sup>38</sup> These moments are summarized in Table 2. Note, in particular, that the model replicates the positive correlations observed in the data between sticker tuition, net tuition, family income and ability: higher tuition colleges tend to enroll students from more affluent backgrounds and higher-ability students. The model slightly over-predicts dispersion in average family income.<sup>39</sup>

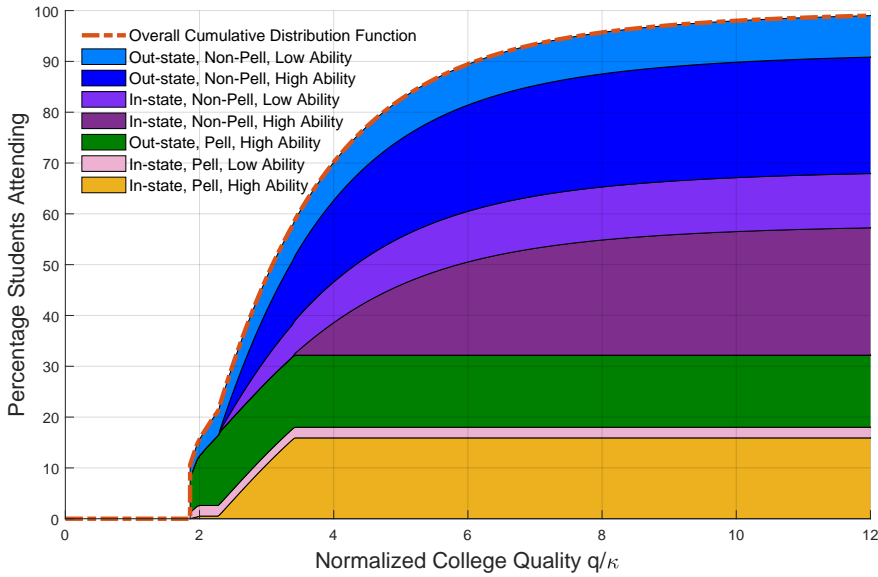
Overall, our relatively simple calibrated model successfully replicates some important features of the US college market: the number of people who attend, attendees' average income and ability, and observed variation in college tuition. We now explore and attempt to better interpret the equilibrium patterns of college enrollment and the nature of equilibrium college pricing.

Figure 5 plots the cumulative distribution function (CDF) for student enrollment under the baseline calibration, conditional on enrolling. The different colors within the dis-

<sup>38</sup>We have no information at the college level about AFQT scores, which we used to measure the share of high ability students enrolled in college. The College Scorecard does report various percentiles of SAT and ACT scores for admitted students. We assume these scores are normally distributed with college-specific means and variances, estimate the fraction of students whose score lies in the top half of the national test score distribution, and identify this fraction as the share of the college student body that is high ability. Another measurement issue is that in the College Scorecard data, 26% colleges do not report information on test scores.

<sup>39</sup>The model also underpredicts dispersion in the fraction of high ability students. However, as noted above in footnote 38, our empirical measure for the share of high ability students is likely quite noisy, which may account for this discrepancy.

Figure 5: Cumulative Quality Distribution and Decomposition

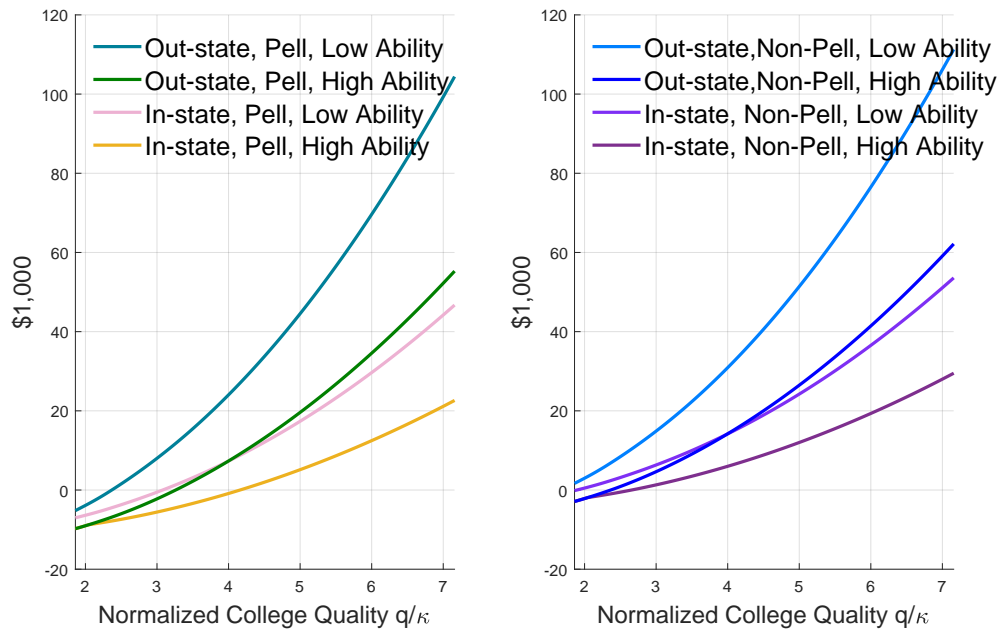


tribution reflect different types of students, where each type faces a type-specific net tuition schedule. There are eight different student types in the model, corresponding to each possible combination of residence status (determining eligibility for in-state tuition discounts), family income relative to the threshold  $y^*$  (determining eligibility for Pell Grants and other need-based aid), and ability (determining eligibility for institutional aid).

Several features of the plot are worth noting. First, no students attend very low-quality colleges: demand for such colleges is weak, given the lost-earnings cost  $\omega$  of attending college in combination with a positive reservation quality  $\kappa$  when not attending. Second, need-based aid eligible students are clustered in relatively low quality colleges (the corresponding conditional CDF's flatten out around normalized quality equal to 3). This suggests strong sorting by income in college enrollment, in line with [Belley and Lochner \(2007\)](#) and [Chetty et al. \(2020\)](#).

To better understand quality choices by student type, [Figure 6](#) plots net tuition by type as a function of quality, while [Figure 7](#) plots the corresponding college quality decisions. Tuition is increasing in quality, in a non-linear fashion. It is lowest for high ability in-state students (who receive both in-state subsidies and merit-based institutional aid) and highest for low-ability out-of-state students (who benefit from neither form of tuition discount). Net tuition for those eligible for need-based aid is simply reduced by the amount of that aid,  $p_1 = \$6,869$ . Note also that tuition for high-ability students rises more slowly

Figure 6: Net Tuition Schedules for Different Types



with quality than tuition for low-ability students.<sup>40</sup>

Each panel of Figure 7 plots college enrollment choices for households of a different combination of ability and residence status. All high-ability students with family income above the need based aid eligibility threshold  $y^*$  enroll in college, as do some poorer households. For low ability students, who face higher tuition, quality choices are non-zero only at higher income values, indicating that a smaller fraction of this group attends college. Note that some low ability in-state students with family income just below  $y^*$  enroll (the disconnected blue line segment), even though similar students from families with income just above  $y^*$  do not.

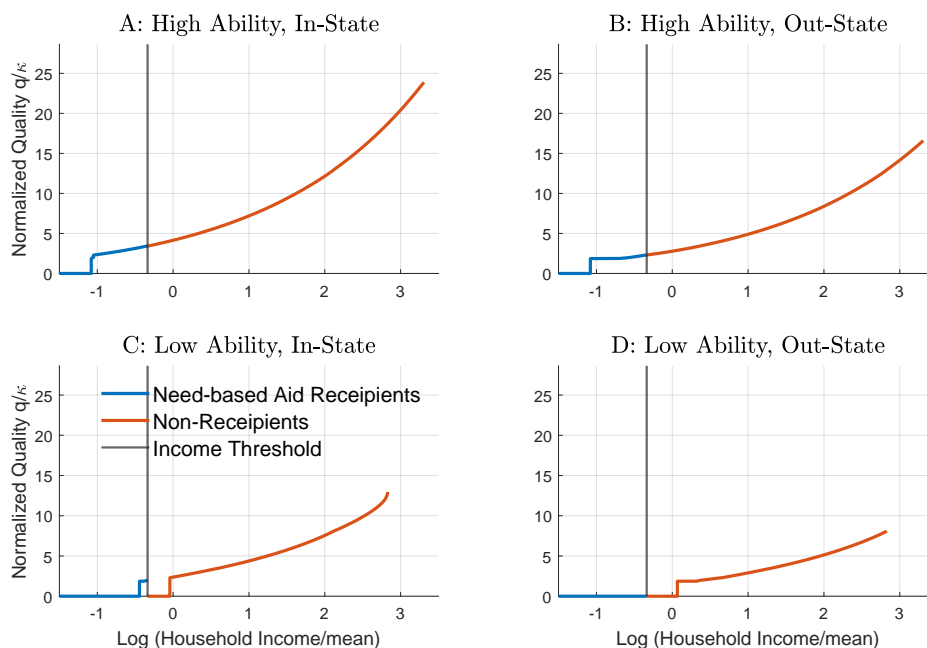
Conditional on enrolling, quality choices are generally increasing in income for a given household type.<sup>41</sup> In addition, because high-ability students face flatter tuition schedules, they choose higher quality colleges than otherwise identical low-ability students.

Figure 7 highlights that the effects on college enrollment and quality choices of different forms of tuition discount are quite different. Need-based aid incentivizes some

<sup>40</sup>Note that net tuition is negative for some household types at low-quality colleges, reflecting the fact that the combination of general, need-based, and merit-based aid can exceed sticker tuition. However, the total cost of attending college, inclusive of the \$10,020 opportunity cost of work, is positive for all types.

<sup>41</sup>Note that a student from a household with income slightly below  $y^*$  who is attending college is effectively slightly richer than a household slightly above  $y^*$ , since the former receives need-based aid  $p_1$ . As a result, the former chooses a slightly higher quality college.

Figure 7: Student Allocation across Colleges



low-income students to enroll in college, especially those with high ability. But because the amount of need-based aid is fixed, this form of subsidy does not incentivize choosing a higher quality college, conditional on enrolling: the quality choices of students just below and just above  $y^*$  in Figure 7 are similar. Thus, in Figure 5 students receiving government need-based aid tend to be clustered in relatively low quality colleges.<sup>42</sup>

In contrast, in-state-resident subsidies in the model do little to encourage enrollment, because at low quality colleges, tuition is already low and proportional in-state discounts are therefore small. Thus, the share of in-state students enrolling in college is similar to their population share: 54.8 percent versus 53.2 percent. However, because in-state discounts are proportional to tuition, the marginal cost of attending higher quality colleges is lower for in-state students, which incentivizes them to choose higher quality colleges (see Figure 7).

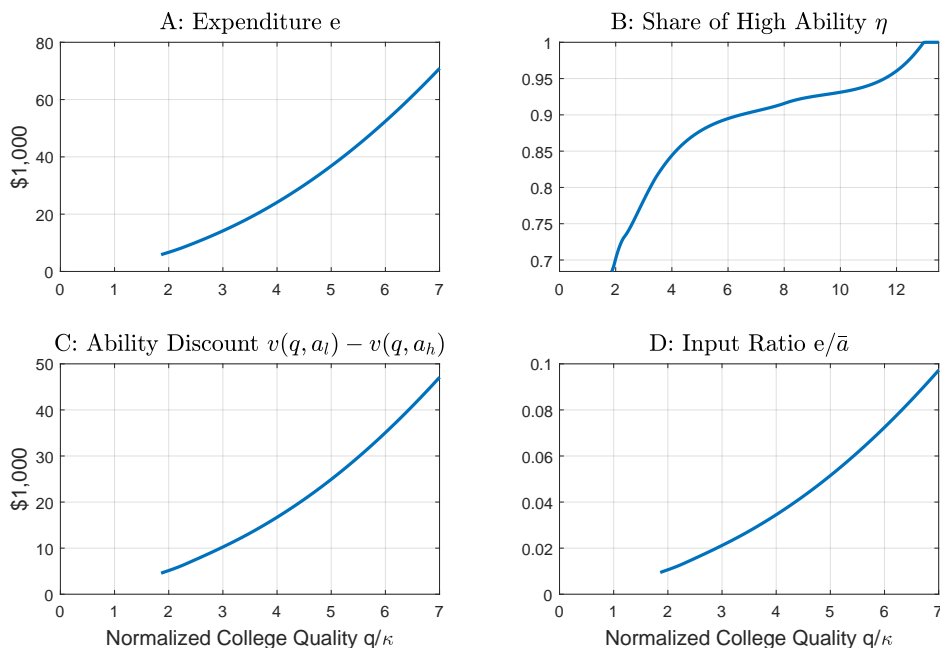
As noted above, the endogenous tuition discounts that high ability students enjoy (institutional aid) incentivize both higher enrollment and higher quality choices, because high ability students face both lower and flatter tuition schedules.<sup>43</sup>

<sup>42</sup>This is also a qualitative feature of the data: 64.2 percent of students at for-profit colleges are Pell Grant recipients (Kelly et al., 2019), compared with around 15 percent at Ivy League universities (see <https://www.romanhighered.com/data-viz/2018/1/2/pell-grant-recipients-in-the-ivy-league>).

<sup>43</sup>Higher enrollment also reflects the fact that high-ability students tend to come from more affluent families.



Figure 8: College Inputs

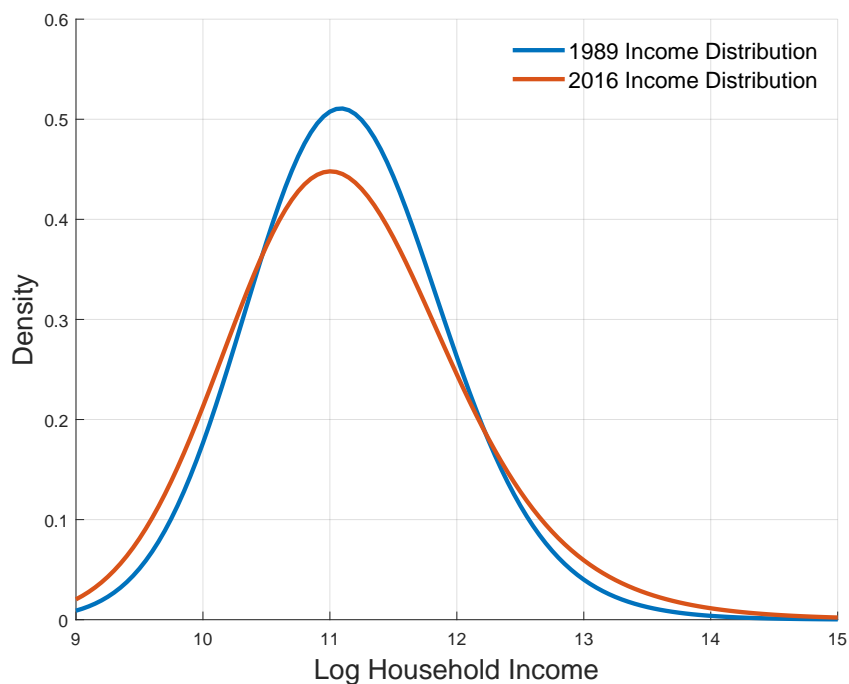


We now turn to the supply side of the market. Figure 8 shows how college inputs vary by quality. Panels A and B plot the level of expenditure per student by quality and the fraction of high-ability students enrolled. Higher quality schools spend more per student and also generally admit a student pool with higher average ability. Panel C indicates that the tuition discount per unit of ability  $d(q)$  (equivalently, the equilibrium differential between tuition for high- and low-ability students) is increasing in college quality. Thus, as quality increases, the effective price of the ability input rises, inducing colleges to increase the ratio of expenditure to average student ability (Panel D; see also eq. 8). Because the ratio  $e/\bar{a}$  increases in quality, the ratio of expenditure to quality is also increasing in quality (from eq. 2, this ratio is  $(e/\bar{a})^{1-\theta}$ ). This convexity in the equilibrium expenditure level (Panel A) translates into convexity in the equilibrium tuition schedules in Figure 6. Intuitively, because high ability students command ever larger tuition discounts moving up the quality distribution, colleges choose instead to spend ever larger amounts on instructional inputs and pass those expenses on in the form of higher tuition.

## 5 Understanding Changes in College Tuition

We now explore possible drivers of the increase in college tuition over time. Our focus is on the role of widening income inequality. We re-estimate the income distribution

Figure 9: Estimated EMG Distribution of Log Household Income (SCF)



parameters  $\sigma^2$  and  $\alpha$  using household income data from the 1989 wave of the Survey of Consumer Finances. We find that  $\sigma^2 = 0.48$  and  $\alpha = 2.4$ . Thus, the implied variance of log income in 1989 is  $0.48 + 2.4^{-2} = 0.65$ , compared with 0.91 in 2016. Note that most of the increase in the variance over this period is attributable to a heavier right tail in the income distribution.<sup>44</sup> Figure 9 plots the estimated exponentially modified Gaussian (EMG) distribution for log household income in the two years. Note that while mean *log* income is very similar in both years, mean *level* income is 24.7 percent larger in 2016.

Another set of changes we can measure and explore are those connected to changes in public support for higher education. Assuming no change in the fixed cost of college education per student  $\phi$ , we can measure the change in total public support per student between 1990 and 2016 as the change in average educational expenditure per student minus the change in average net tuition.<sup>45</sup> We estimate average educational spending in

<sup>44</sup>Data from the online appendix to Piketty and Saez (2003) indicate a similar increase in the Pareto parameter. For example, one can estimate the Pareto parameter as average income conditional on being above the  $x^{th}$  percentile of the income distribution, relative to this average minus income at the  $x^{th}$  percentile. Applying this formula to the Piketty-Saez data at the 90<sup>th</sup> percentile of the income distribution implies Pareto coefficients of 2.20 in 1989 and 1.83 in 2014.

<sup>45</sup>While the difference between college spending and net tuition paid primarily reflects various forms of government subsidy, it also captures other sources of non-tuition revenue such as endowment income or

Table 3: Summary Changes in College Market, 2016 versus 1990

	2016 Data	1990 Data	% Growth
Net tuition	\$9,250	\$6,034	53.3
Expenditure per student $e$	\$17,077	\$10,503	62.6
Total subsidies per student net of fixed cost $\phi$	\$7,828	\$4,469	75.2
Need-based aid	\$2,198	\$1,377	59.6
In-state subsidies	\$8,343	\$5,413	54.1
General subsidies to colleges net of $\phi$	-\$4,609	-\$2,396	-16.9
General subsidies to students	\$1,896	\$76	
Enrollment	0.507	0.327	+18.0pp
Share in-state	0.546	0.581	-3.5pp
Share Pell	0.32	0.30	+2.0pp
Graduation	0.361	0.233	+12.8pp

1990 to be \$10,503, compared with \$17,077 in 2016. We estimate average net tuition in 1990 to be \$6,034, compared with \$9,250 in 2016. The fact that spending has increased much more than out-of-pocket tuition indicates growth of \$3,359 in per student public subsidies.

In Table 3, we decompose this \$3,359 growth in public support into (1) growth in public need-based aid, (2) growth in public subsidies to in-state students, and (3) growth in general per student public subsidies. We estimate the values of these three components, on a per student basis, to be \$821, \$2,930, and  $-\$393$ , respectively. Thus, need-based aid has risen, the value of in-state subsidies has risen quite markedly, and the value of general subsidies has declined slightly. The composition of these general subsidies has also changed, according to our estimates. In particular, general subsidies to colleges,  $\bar{s}$ , have *declined* by \$2,213, while general subsidies to students,  $p_0$ , have *risen* by \$1,820. In the context of our competitive model, whether colleges or students receive per student subsidies is irrelevant for the equilibrium allocation and for net tuition, but a shift in subsidies towards students implies a commensurate increase in sticker tuition.<sup>46</sup>

## 5.1 The Effect of Rising Income Inequality

We first explore the implications of changes in income inequality for college enrollment and tuition. Column (2) of Table 4 reports how this change in inequality changes some key

alumni giving. Thus, our “public support” label includes these sources of income, in addition to government subsidies.

<sup>46</sup>We do not experiment with changing the opportunity cost of attending college  $\omega$ , since inflation-adjusted median earnings for workers between the ages of 16 and 24 barely changed between 1990 and 2016.

Table 4: Effects of Changes in Income Inequality and Subsidies

	(1)	(2)	(3)	(4)	(5)
	2016 Model	1989 Ineq.	1989 Mean	1989 Dist.	Dist.+ Subs.
<i>Parameters changed</i>	–	$\sigma^2, \alpha$	$\bar{y}$	$\bar{y}, \sigma^2, \alpha$	all in Table 5
Net tuition	\$9,250	\$7,359	\$7,746	\$5,921	\$7,476
Expenditure	\$17,077	\$13,904	\$13,982	\$10,944	\$12,389
Enrollment	0.507	0.562	0.428	0.477	0.447
Family inc. enrolled / mean	1.567	1.407	1.747	1.543	1.625
Share high ability enrolled	0.802	0.889	0.714	0.807	0.746
Share low ability enrolled	0.212	0.235	0.143	0.147	0.149
Quality / $\kappa$	3.724	3.604	3.569	3.448	3.594

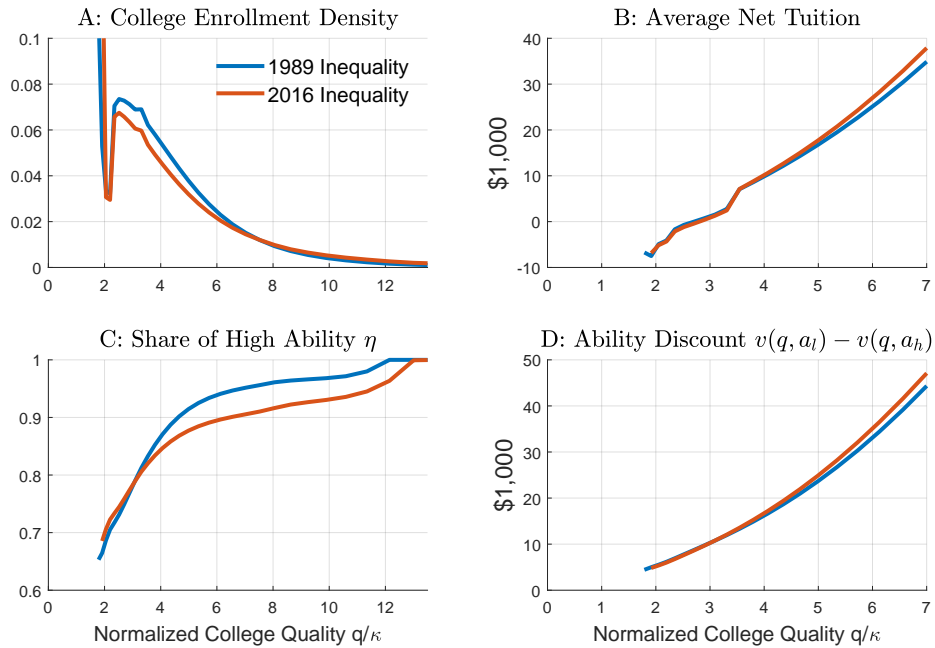
Table 5: Changes in Income Distribution and Subsidy Parameters

2016 Income distribution		1989 Income distribution	
$\bar{y}$	\$774,590	$\bar{y}$	\$621,221
$\sigma^2$	0.548	$\sigma^2$	0.478
$\alpha$	1.67	$\alpha$	2.40
2016 Subsidies		1990 Subsidies	
$\mu_i$	0.529	$\mu_i$	0.578
$\lambda$	0.490	$\lambda$	0.536
$y^* / \bar{y}$	0.714	$y^* / \bar{y}$	0.884
$p_1$	\$6,870	$p_1$	\$4,590
$p_0$	\$1,896	$p_0$	\$76
$\phi - \bar{s}$	\$4,610	$\phi - \bar{s}$	\$1,954

predictions of the model. In order to isolate the impact of changing income dispersion, the mean parameter  $\mu$  is re-scaled so that average household income  $\bar{y}$  for the economy in column (2) is identical to the one in column (1).

The table indicates that the observed change in income inequality over this period predicts large changes in net tuition and enrollment. Given income inequality in 1989, average net tuition in the model is \$7,359, suggesting that changing income inequality alone can account for over half of the observed rise in net tuition over this period (\$1,891 out of \$3,216, or 58.8%). Average net tuition rises because households at the top of the income distribution became much richer and thus more willing to pay for expensive high quality colleges. At the same time, the counterpart to fast income growth at the top of the distribution was declining relative income for households in the middle of the distribution. These households were close to indifferent about attending college in 1989, and these income losses therefore drive down college enrollment. Thus, the model pre-

Figure 10: Effect of Changing Income Inequality



dicts that rising income inequality (all else held equal) reduced the enrollment rate by 5.5 percentage points.

Figure 10 illustrates how changing income inequality affects the equilibrium quality distribution (Panel A), the share of high ability students by quality (Panel C), and equilibrium tuition schedule (Panels B and D). Increasing income inequality reduces the equilibrium supply of low quality schools while increasing slightly the number of spots at very high quality schools. This helps to explain why average net tuition rises. Across most of the quality distribution, the share of high-ability students in college declines, indicating that colleges rely more on instructional spending and less on peer effects to maintain quality.

One might suspect that rising income inequality would lead to college attendance's being driven more by income and less by ability. Indeed, when inequality is increased, the share of high ability students enrolling in college falls by 8.7 percentage points, as poorer high ability students are effectively priced out. At the same time, however, increasing inequality also increases the *demand* for high-ability students. In particular, because the quality production technology features decreasing returns to expenditure, colleges seeking to satisfy increased demand for high quality want to increase both expenditures and average ability, all else equal. Thus, increased demand for high quality colleges indirectly

increases the relative demand for high ability students, which translates into larger institutional tuition discounts for high-ability students (see Panel D of Figure 10). This suggests that rising income inequality has played a role in generating the observed growth in institutional financial aid and the associated rising gap between sticker and net tuition (see Figure 1). Note that the increasing price of attracting high ability students (Panel D of Figure 10) explains why colleges at each quality level rely more on instructional expenditure and less on a high ability student body when inequality rises (Panel C).

To recap, when inequality rises both net tuition and expenditure rise as colleges respond to increased demand from households at the top of the income distribution for high quality. Moving from column (2) to column (1), net tuition grows by 25.7 percent and instructional expenditure per student rises 22.8 percent, but this extra spending leads to only a 3.3 percent increase in average quality. The increase in average quality is small for two reasons.

The first is the presence of peer effects: expenditure is only one quality-enhancing input, and the other — average peer ability — does not increase in response to rising income inequality. If quality were equal to expenditure, as in a conventional non-club good model, then average quality would obviously rise by exactly as much as expenditure per student.

The second reason the increase in average quality is small is that the increase in expenditure is concentrated in colleges at the top of the quality distribution, where expenditure is already high and the marginal product of additional spending is low. Given equal factor shares ( $\theta = 0.5$ ), our production technology implies that average college quality is given by

$$E[q] = E[\bar{a}(q)^{0.5}e(q)^{0.5}] = \sqrt{E[\bar{a}] - \text{var}(\sqrt{\bar{a}})}\sqrt{E[e] - \text{var}(\sqrt{e})} + \text{cov}(\sqrt{\bar{a}}, \sqrt{e}).$$

When income inequality is increased, the boost to average quality from higher average expenditure is partially offset by a rise in the variance of (the square root of) expenditure, coupled with a fall in the correlation between average ability and expenditure. If we were to observe the same changes in average ability and expenditure in a model with only one type of college (so that  $E[q] = \sqrt{E[\bar{a}]} \sqrt{E[e]}$ ), then average quality would increase by 10.8 percent.

We conclude that in order to properly quantify the rise in average college quality associated with a rise in income inequality, it is important both to model peer effects and to allow for heterogeneity in college quality: abstracting from either feature leads one

to overstate the increase in average quality and thus to understate growth in quality-adjusted tuition.

## 5.2 The Roles of Other Drivers of Tuition Growth

Column (3) of Table 4 shows the effect of reducing average income (via a reduction in the parameter  $\mu$ ) to its value in 1989.<sup>47</sup> Reducing average income reduces households' willingness to pay for college and thus lowers both the enrollment rate and net tuition. Thus, the model points to economic growth as one factor driving up college tuition.

Column (4) shows the combined effect of reducing both income inequality and mean income to their values in 1989. In this case, the model implies average net tuition of \$5,921, which is very similar to the actual value in 1990. The predicted enrollment rate given the income distribution in 1989 is lower than in 2016, indicating that the positive effect on enrollment of growth in average income outweighs the negative effect of widening inequality.

Column (5) of Table 4 shows how key model predictions change when, in addition to moving back to the income distribution in 1989, we also roll back subsidies to their values in 1990. In particular, we change the six subsidy parameters to the values described in Table 5. The values in 1990 for these parameters are chosen to match the four subsidy values in 1990 in Table 3 as well as the in-state and Pell-eligible enrollment shares for 1990 listed there.<sup>48</sup>

Lower subsidies per student in 1990 translate into higher tuition and a lower enrollment rate relative to the economy with subsidies at the 2016 values (compare column 5 with column 4). Thus, we find that growth in public subsidies (and other non-tuition revenue) has played a role in supporting enrollment and restraining growth in tuition.

We have also experimented with changing one aspect of the technology for producing college quality—namely the price, relative to consumption, of the variable instructional input,  $e$ . By allowing this price to rise over time, we can assess whether rising faculty salaries play a significant role in accounting for rising tuition (see, e.g., [Jones and Yang](#)

---

<sup>47</sup>In this experiment, we also reduce the need-based aid threshold  $y^*$  proportionately with mean income. No other parameters are changed relative to the economy in column (1).

<sup>48</sup>Need-based aid in 1990 is straightforward to estimate, given 1990 data on average Pell Grants and need-based state grants. Conditional on receiving aid, need-based aid has risen from \$4,589 to \$6,870, or from \$1,376 to \$2,198 on an unconditional per student basis.

Growth in subsidies to in-state students is more difficult to estimate. In 2016, we measured the in-state subsidy parameter  $\lambda$ , using the ratio of in-state to out-of-state sticker tuition for public colleges. We were unable to find direct estimates of out-of-state public college tuition in 1990. We therefore impute a value for out-of-state public sticker tuition in 1990 by multiplying private sticker tuition in 1990 by the 2016 ratio of public out-of-state sticker tuition to private sticker tuition.

2016). We estimate that this price rose by 9.4 percent in real terms over the 1990–2016 period.<sup>49</sup> However, reducing the cost of the instructional input has a small impact on average net tuition or the enrollment rate.<sup>50</sup> The logic is that when college quality becomes more expensive to produce, for the most part households simply reduce their chosen college quality rather than increase the fraction of income they devote to college.

Overall, these experiments suggest that rising income inequality plays the most important role in accounting for observed growth in college tuition, while rising average income is also an important factor. Rising subsidies (and non-tuition revenue more generally) have a modest offsetting impact in the opposite direction. Rising costs of instructional inputs play a marginal role.

### 5.3 Changes in Enrollment Patterns

Comparing across all the different experiments, the only parameter change that has a notable impact on enrollment of low ability students is growth in average income: the enrollment share of low-ability students is 6.9 percentage points larger in the baseline calibration for 2016 compared with when average income is reduced to its value in 1989 (column 3). For high-ability students, in contrast, all the different factors considered affect enrollment. Growth in average income raises high-ability enrollment by 8.8 percentage points (column 1 versus 3), increasing inequality reduces high-ability enrollment by 8.7 to 9.3 percentage points (column 1 versus 2, or 3 versus 4), and increasing subsidies increase enrollment by 6.1 percentage points (column 4 versus 5). Thus, while rising income inequality has reduced enrollment by high-ability students from poor families, the effect has been partially offset by growth in need-based financial aid.

Overall, the combined effects of changes in the income distribution and in subsidies are to increase the share of high-ability students enrolling in college by 5.6 percentage points and to increase the share of low-ability students enrolling by 6.3 percentage points (comparing columns 1 and 5). All of these extra low-ability students are drawn from the top half of the model family income distribution.

[Belley and Lochner \(2007\)](#) describe patterns of college attendance by ability (AFQT score) and family income for college-age individuals in the 1979 and 1997 NLSY.<sup>51</sup> They

---

<sup>49</sup>Formally, we assume that  $z_t$  units of the consumption good  $y$  are required to produce one unit of the instructional input  $e$ . Given a competitive market for  $e$ , the equilibrium price of  $e$  at date  $t$  is then  $p_t = z_t$ . We normalize  $z_{2016} = 1$ .

<sup>50</sup>Reducing the input price raises enrollment by 1.5 percentage points, reduces net tuition by \$348, and increases average college quality by 3.61%.

<sup>51</sup>Note that our quantitative exercise focuses on a comparison between 1990 and 2016, while the two



Table 6: The Role of Peer Effects

	$\theta = 0.25$	$\theta = 0.5$	$\theta = 0.75$
Enrollment Pattern			
Family income enrolled / mean	2.020	1.567	1.498
Share of high ability enrolled	0.731	0.802	0.876
Share of low ability enrolled	0.283	0.212	0.139
Impact of Rising Inequality (1989 to 2016)			
Enrollment rate (change, percentage points)	- 6.30	- 5.46	- 3.74
Net tuition (change, \$)	+ 453	+1,891	+2,210

find that almost half of the increase in college attendance between the two waves was from students in the bottom half of the AFQT score distribution. Furthermore, these additional low-ability college students were drawn mainly from the top half of the income distribution. Thus, both model and data are consistent with the message that income has become a more important driver of college attendance, relative to student ability.

#### 5.4 The Importance of Peer Effects

How important are peer effects for our results? To address this question we have experimented with alternative values for  $\theta$ , which governs the relative importance of peers versus expenditure in producing college quality. We experimented with setting  $\theta = 0.25$  and  $\theta = 0.75$ . In each case, we recalibrated all model parameters, following our baseline calibration strategy.<sup>52</sup>

When peer effects are less important ( $\theta = 0.25$ ), we find that income becomes a more important driver of college enrollment: the average family income of those enrolling in college is now 2.02 times the average, compared with 1.57 in the baseline model (and 1.56 in the data). The intuition is simple: when quality depends mostly on resources, high-ability students enjoy smaller tuition discounts, and income becomes more important, relative to ability, in driving the enrollment pattern. When peer effects are more important ( $\theta = 0.75$ ), in contrast, the model predicts that ability becomes more important in driving college enrollment: 88 percent of high ability students now enroll in college. The fact that the model generates a pattern of college enrollment by income and ability similar to the one observed in the data when  $\theta = 0.5$  offers indirect support for our baseline parameter choice.

NLSY waves offer a comparison of the early 1980s with the early 2000s.

<sup>52</sup>We hold the values for  $a_h$  and  $a_l$  fixed at their baseline values across these experiments, because changing those values also influences the importance of peer effects.

We also revisited our key experiment of changing income inequality for these alternative parameterizations. What we find is that with a larger value for  $\theta$ , the model delivers a larger increase in average net tuition in response to an increase in income inequality, and a smaller decline in aggregate enrollment. These results are summarized in Table 6. The finding that as peer effects become more important, changes in inequality have a smaller impact on equilibrium quantities and a larger impact on prices is qualitatively consistent with our earlier analytical result from the simpler model described in Section 3, where we showed that in the limiting case in which quality is a pure peer effect, changing inequality has zero impact on the allocation of quality and affects only equilibrium tuition.

## 6 Education as Investment

In this section, we briefly discuss an alternative setting in which education is treated as an investment good instead of a consumption good. We will contrast the nature of equilibrium allocations in these two settings and the role of access to a credit market in the investment model.

In the investment model, parents do not directly care about their childrens' education, but instead they care about their future income and consumption, which depends, in part, on education. Parents consume, borrow, or lend, and potentially invest in college in an initial period. Children generate income and consume and repay loans in a second period. For expositional purposes, we assume away government policies such as in-state subsidies and need-based aid and also abstract from drop-out risk.<sup>53</sup>

Credit is in zero net supply, and households borrow and lend at an equilibrium interest rate  $R$ . Loans must be repaid, but to start, there are no additional restrictions on credit. The tuition function is given by  $t(q, a)$ . Taking the market interest rate and the tuition function as given, the household's problem is

$$\begin{aligned} & \max_{c_1, c_2, q, b} \{ \log(c_1) + \beta \log(c_2) \} \\ & \quad s.t. \\ & c_1 + t(q, a) \leq y + b - \mathbb{1}_{\{q > 0\}} \omega \\ & c_2 = y_2(q, a) - Rb, \end{aligned}$$

---

<sup>53</sup>Incorporating drop-out risk in an investment setup would require a discussion of whether this risk can be insured.

where the second period income function is

$$y_2(q, a) = (q + \tau)^\zeta a^\lambda, \text{ where } \tau, \zeta, \lambda > 0. \quad (10)$$

The interpretation is that the child's income depends on the quality of her education  $q$  and on her ability  $a$ , with corresponding elasticities defined by  $\zeta$  and  $\lambda$ . The solution to this problem is a consumption rule  $c_1(y, a), c_2(y, a)$ , an education choice  $q(y, a)$ , and a borrowing decision  $b(y, a)$ .

The college aspect of the model remains unchanged, with college decision rules  $e(q)$  and  $\eta_a(q)$  and an endogenous distribution of college quality  $\chi(q)$ . A competitive equilibrium can be defined similarly to the baseline college-as-consumption model, with one extra market-clearing condition for the credit market.

**Proposition 6.** *The competitive equilibrium in the investment model with frictionless credit is Pareto efficient.*

**Proposition 7.** *The competitive equilibrium of the investment model with frictionless credit features no sorting by income in college enrollment and positive sorting by ability.*

The intuition is that given the ability-education complementarity (equation 10), it is efficient to admit high ability students to college, regardless of their income. Since the competitive equilibrium must be efficient (Proposition 6), the equilibrium enrollment pattern features sorting only by ability. Note that the lack of sorting by income in the frictionless investment model is at odds with the empirical patterns documented in the data by [Belley and Lochner \(2007\)](#) and others, especially for recent decades.

We now introduce frictions into the credit market and show that as frictions increase, this model converges to our college-as-consumption model, which features positive sorting by income.

## 6.1 Frictional Credit Market

We now introduce an exogenous wedge  $\rho$  between the borrowing rate and the lending rate, which captures frictions in the credit market. The lending rate is  $R$  and the borrowing rate is  $R + \rho$ . As  $\rho \rightarrow 0$ , the model collapses to the frictionless case just described.<sup>54</sup>

---

<sup>54</sup>This is a highly stylized model. Embedding our model of the college market into a life-cycle framework with a quantitative model of the student loan market is a possible avenue for future research (see, for example, [Abbott et al. 2019](#)).

The households' problem is identical to the one just described, except that the second period budget constraint is now

$$c_2 = y_2(q, a) - \left\{ R + \mathbb{1}_{\{b>0\}}\rho \right\} b.$$

With a positive wedge  $\rho > 0$ , the equilibrium is not efficient. Intuitively, poor families with high ability children may under-invest in education when credit is costly.

When the credit friction is sufficiently severe ( $\rho$  is sufficiently large) the frictional investment model collapses to the consumption model we used for our quantitative analysis. The intuition is that when the credit friction is sufficiently severe, there is no credit in equilibrium:  $b(y, a) = 0$  for any pair  $(y, a)$ . Substituting the implied expression for  $c_2$  into the objective function, the household's problem is then isomorphic to the one in the college-as-consumption model:

**Proposition 8.** *There exists a threshold  $\bar{\rho}$  such that for any  $\rho > \bar{\rho}$ , the investment model is isomorphic to a consumption model in which  $\varphi = \beta\zeta$  and  $\kappa = \tau$ .*

Thus, the consumption model considered in the main text can be reinterpreted as an investment model in which the financial friction is severe. In both these models, income plays an important role in college attendance. The frictionless investment model instead implies a quite different pattern, in which enrollment is independent of income. That implication is inconsistent with the data, while our baseline college-as-consumption model delivers a realistic pattern of college enrollment across the income-ability distribution. For that reason, we have focused on the college-as-consumption model in our quantitative application.

## 7 Conclusion

A satisfactory model of the college market is essential for understanding what sorts of potential students go to college, what sorts of colleges they attend, and how much they pay. It is also important for understanding how these features of the college landscape have changed over time, and for exploring the impact of possible policy interventions.

We have developed a competitive model of the college market in which college quality depends on the average ability of attending students. A novel feature of the model is a continuous distribution of college quality. When we use a calibrated version of the model to predict the impact of the rise in top-tail income inequality since 1989, we find that greater income inequality can explain over half of the observed increase in average net

tuition. By itself, greater income inequality would depress enrollment rates, but we find that growth in average income and more generous college subsidies are countervailing forces that have pushed more people into college.

Our analysis could be extended in many directions. First, while a perfectly competitive model offers a reasonable positive theory of observed outcomes, competition is likely not perfect in practice. [Fillmore \(2016\)](#) argues that colleges exploit information on FAFSA forms to price discriminate by income. He shows that higher family income translates into smaller tuition discount offers, even after controlling for observable proxies for ability (ACT scores and high school GPAs). This pattern emerges primarily at the most selective schools, suggesting that such colleges have some pricing power (for example, see [Epple et al. 2017, 2019](#)). One could allow colleges to earn rents (and thereby consider alternative objectives to profit maximization) by endowing colleges with an idiosyncratic non-reproducible attribute, such as location or brand name, and endowing households with heterogeneous preferences over this attribute.

A second possible extension would be to model heterogeneity in college information about student ability. At the time of admission, students would then be unable to perfectly anticipate the terms of admission offers and would therefore want to apply to multiple schools (see, e.g., [Fu 2014](#)).

A third avenue for further research is an exploration of the nature of optimal college subsidies, tying subsidies to student attributes (income or ability), the quality of the college the student attends, or both (see [Findeisen et al. 2018](#)). The nature of the optimal intervention will depend on the planner's social welfare function, in addition to whether the laissez-faire allocation is efficient (as when college is a pure consumption good) or inefficient.

Fourth, one could consider broader sets of attributes that may affect students' attractiveness as peers, in addition to measures of performance on standardized tests. For example, one way to introduce income-dependent pricing in a competitive framework would be to posit that students perceive value to having peers from diverse backgrounds, which would make minority or low-income students attractive to colleges that draw predominantly from rich white families.<sup>55</sup>

A final possible application would be to develop a multi-generational extension of the model outlined in Section 6 to explore the propagation of inequality across generations. Consider an increase in the financial return to college quality. This will lead to an increase

---

<sup>55</sup>For example, one could define peer desirability (ability) as  $a = s^\kappa y^{1-\kappa}$ , with  $s$  being a test score measure and  $y$  being family income. Given  $\kappa > 1$ , our competitive framework would then imply larger tuition discounts for low family income students, all else equal.

in investment in quality by higher-income households, which will amplify the effect on income inequality in the next generation. In turn, a fatter right tail in the income distribution for that generation will further amplify inequality in college investment. Over successive generations, a small increase in the return to college quality can potentially generate both a large increase in income inequality and a decline in intergenerational mobility.

## References

- Abbott, Brant, Giovanni Gallipoli, Costas Meghir, and Giovanni L. Violante (2019), "Education Policy and Intergenerational Transfers in Equilibrium," *Journal of Political Economy*, 127(6), 2569–2624.
- Abdulkadiroğlu, Atila, Joshua Angrist, and Parag Pathak (2014), "The Elite Illusion: Achievement Effects at Boston and New York Exam Schools," *Econometrica*, 82(1), 137–196.
- Angrist, Joshua D. (2014), "The Perils of Peer Effects," *Labour Economics*, 30, 98–108.
- Arnott, Richard, and John Rowse (1987), "Peer Group Effects and Educational Attainment," *Journal of Public Economics*, 32(3), 287–305.
- Belley, Philippe, and Lance Lochner (2007), "The Changing Role of Family Income and Ability in Determining Educational Achievement," *Journal of Human Capital*, 1(1), 37–89.
- Black, Sandra E. (1999), "Do Better Schools Matter? Parental Valuation of Elementary Education," *The Quarterly Journal of Economics*, 114(2), 577–599.
- Boustan, Leah Platt (2012), "School Desegregation and Urban Change: Evidence from City Boundaries," *American Economic Journal: Applied Economics*, 4(1), 85–108.
- Carrell, Scott E., Richard L. Fullerton, and James E. West (2009), "Does Your Cohort Matter? Measuring Peer Effects in College Achievement," *Journal of Labor Economics*, 27(3), 439–464.
- Caucutt, Elizabeth M. (2001), "Peer Group Effects in Applied General Equilibrium," *Economic Theory*, 17(1), 25–51.
- Caucutt, Elizabeth M. (2002), "Educational Vouchers When There Are Peer Group Effects—Size Matters," *International Economic Review*, 43(1), 195–222.
- Chetty, Raj, John N. Friedman, Emmanuel Saez, Nicholas Turner, and Danny Yagan (2017), "Mobility Report Cards: The Role of Colleges in Intergenerational Mobility," Working Paper 23618, National Bureau of Economic Research.
- Chetty, Raj, John N. Friedman, Emmanuel Saez, Nicholas Turner, and Danny Yagan (2020), "Income Segregation and Intergenerational Mobility Across Colleges in the United States," *The Quarterly Journal of Economics*, 135(3), 1567–1633.
- Chetty, Raj, Nathaniel Hendren, Patrick Kline, and Emmanuel Saez (2014), "Where Is the Land of Opportunity? The Geography of Intergenerational Mobility in the United States," *Quarterly Journal of Economics*, 129(4), 1553–1623.
- Cole, Harold L., and Edward C. Prescott (1997), "Valuation Equilibrium with Clubs," *Journal of Economic Theory*, 74(1), 19–39.
- Coleman, James S. (1966), "Equality of Educational Opportunity (COLEMAN) Study (EEOS)," *Inter-University Consortium for Political and Social Research*.

- College Board (2016), *Trends in College Pricing 2016*, trends in Higher Education Series. Available at <https://trends.collegeboard.org>.
- Dale, Stacy B., and Alan B. Krueger (2014), "Estimating the Effects of College Characteristics over the Career Using Administrative Earnings Data," *Journal of Human Resources*, 49(2), 323–358.
- Ellickson, Bryan, Birgit Grodal, Suzanne Scotchmer, and William R. Zame (1999), "Clubs and the Market," *Econometrica*, 67(5), 1185–1217.
- Epple, Dennis, and Richard Romano (2008), "Educational Vouchers and Cream Skimming," *International Economic Review*, 49(4), 1395–1435.
- Epple, Dennis, and Richard Romano (2010), "Peer Effects in Education: a Survey of the Theory and Evidence," in *Jess Benhabib, Alberto Bisin, and Matthew O. Jackson (eds.) Handbook of Social Economics*, vol. 1.
- Epple, Dennis, Richard Romano, Sinan Sarpca, and Holger Sieg (2017), "A General Equilibrium Analysis of State and Private Colleges and Access to Higher Education in the U.S." *Journal of Public Economics*, 155(C), 164–178.
- Epple, Dennis, Richard Romano, Sinan Sarpça, Holger Sieg, and Melanie Zaber (2019), "Market Power and Price Discrimination in the US Market for Higher Education," *The RAND Journal of Economics*, 50(1), 201–225.
- Epple, Dennis, Richard Romano, and Holger Sieg (2006), "Admission, Tuition, and Financial Aid Policies in the Market for Higher Education," *Econometrica*, 74(4), 885–928.
- Epple, Dennis, and Richard E Romano (1998), "Competition between Private and Public Schools, Vouchers, and Peer-Group Effects," *American Economic Review*, 88(1), 33–62.
- Fillmore, Ian (2016), "Price Discrimination and Public Policy in the U.S. College Market," Working paper, Washington University.
- Findeisen, Sebastian, Dominik Sachs, and Mark Colas (2018), "Optimal Need-Based Financial Aid," Working paper 14, opportunity and inclusive growth institute, federal reserve bank of minneapolis.
- Fu, Chao (2014), "Equilibrium Tuition, Applications, Admissions, and Enrollment in the College Market," *Journal of Political Economy*, 122(2), 225–281.
- Golden, Daniel (2006), "The Price of Admission: How America's Ruling Class Buys Its Way into Elite Colleges—and Who Gets Left Outside the Gates," *New York: Three Rivers Press*.
- Gordon, Grey, and Aaron Hedlund (2017), "Accounting for the Rise in College Tuition," in *Charles R. Hultan and Valerie Ramey (eds.), Education, Skills, and Technical Change: Implications for Future U.S. GDP Growth*, 357–394 NBER Books.
- Gottschalk, Peter, and Minh Huynh (2010), "Are Earnings Inequality and Mobility Overstated? The Impact of Nonclassical Measurement Error," *The Review of Economics and Statistics*, 92(2), 302–315.
- Heathcote, Jonathan, and Hitoshi Tsujiyama (2019), "Optimal Income Taxation: Mirrlees Meets Ramsey," Working paper, Federal Reserve Bank of Minneapolis.
- Hendricks, Lutz, Tatyana Koreshkova, and Oksana Leukhina (2018), "Sorting of Students into Colleges: Inefficiencies and Policy Implications," Working paper.
- Hoekstra, Mark (2009), "The Effect of Attending the Flagship State University on Earnings: A Discontinuity-Based Approach," *The Review of Economics and Statistics*, 91(4), 717–724.
- Jones, John Bailey, and Fang Yang (2016), "Skill-Biased Technical Change and the Cost of Higher Education," *Journal of Labor Economics*, 34(3), 621–662.

- Kelly, Emily, Laura Holian, and Casey Archer of AnLar (2019), "Trends in Pell Grant Receipt and the Characteristics of Pell Grant Recipients: Selected Years, 2003–04 to 2015–16," Technical report, U.S. Department of Education.
- Lavy, Victor, M. Daniele Paserman, and Analia Schlosser (2012), "Inside the Black Box of Ability Peer Effects: Evidence from Variation in the Proportion of Low Achievers in the Classroom," *The Economic Journal*, 122(559), 208–237.
- Lazear, Edward P. (2001), "Educational Production," *The Quarterly Journal of Economics*, 116(3), 777–803.
- Manski, Charles F. (1993), "Identification of Endogenous Social Effects: The Reflection Problem," *The Review of Economic Studies*, 60(3), 531–542.
- Marmaros, David, and Bruce Sacerdote (2002), "Peer and Social Networks in Job Search," *European Economic Review*, 46(4), 870 – 879.
- Piketty, Thomas, and Emmanuel Saez (2003), "Income Inequality in the United States, 1913–1998," *The Quarterly Journal of Economics*, 118(1), 1–41.
- Rothschild, Michael, and Lawrence J. White (1995), "The Analytics of the Pricing of Higher Education and Other Services in Which the Customers Are Inputs," *Journal of Political Economy*, 103(3), 573–586.
- Sacerdote, Bruce (2001), "Peer Effects with Random Assignment: Results for Dartmouth Roommates," *The Quarterly Journal of Economics*, 116(2), 681–704.
- Sacerdote, Bruce (2014), "Experimental and Quasi-Experimental Analysis of Peer Effects: Two Steps Forward?" *Annual Review of Economics*, 6(1), 253–272.
- Schwartz, Amy Ellen, and Benjamin Scafidi (2004), "What's Happened to the Price of College? Quality-Adjusted Net Price Indexes for Four-Year Colleges," *Journal of Human Resources*, 39(3), 723–745.
- Scotchmer, Suzanne (1997), "On Price-Taking Equilibria in Club Economies with Nonanonymous Crowding," *Journal of Public Economics*, 65(1), 75–88.
- Wai, Jonathan, Matt Brown, and Christopher Chabris (2018), "Using Standardized Test Scores to Include General Cognitive Ability in Education Research and Policy," *Journal of Intelligence*, 6(3), 37.
- Zimmerman, David J. (2003), "Peer Effects in Academic Outcomes: Evidence from a Natural Experiment," *The Review of Economics and Statistics*, 85(1), 9–23.
- Zimmerman, Seth (2019), "Elite Colleges and Upward Mobility to Top Jobs and Top Incomes," *American Economic Review*, 109(1), 1–47.



## 8 Not for Publication Appendices

### 8.1 Theoretical Appendix

#### Proof of Proposition 1

Statement: Consider any equilibrium with an equilibrium tuition schedule  $\tilde{t}(q, y, a, r)$ . The same allocation can be supported by an alternative tuition schedule that is given by

$$t(q, a, r) = v(q, a) - s(q, a, r)$$

where

$$v(q, a) = \max_{y, r} \{ \tilde{t}(q, y, a, r) + s(q, a, r) \}$$

Proof: Fix some  $(q, a)$ . Let  $(y^*, r^*)$  denote a solution to the per student revenue maximization problem

$$(y^*, r^*) \in \arg \max_{y, r} \{ \tilde{t}(q, y, a, r) + s(q, a, r) \}. \quad (11)$$

Suppose that there exists another  $y^{**} \neq y^*$  such that:

$$\tilde{t}(q, y^{**}, a, r^*) < \tilde{t}(q, y^*, a, r^*)$$

Then the demand for students of type  $(y^{**}, a, r^*)$  at colleges of quality  $q$  would be zero as colleges could make strictly higher profit by admitting students of type  $(y^*, a, r^*)$  while maintaining technological feasibility. By market clearing the supply of such students at colleges of quality  $q$  must be zero as well.

Now we can reset the tuition faced by students of type  $(y^{**}, a, r^*)$  at colleges of quality  $q$  to:

$$t(q, y^{**}, a, r^*) = \tilde{t}(q, y^*, a, r^*) > \tilde{t}(q, y^{**}, a, r^*) \quad (12)$$

Given this new tuition value, the supply of students of type  $(y^{**}, a, r^*)$  to colleges of quality  $q$  would still be zero as they now face higher tuition. The demand for such students can still be zero, as the college (weakly) prefers to admit students of type  $(y^*, a, r^*)$ . Thus the market clearing condition still holds. We can continue this procedure for all  $y \neq y^*$  until for all  $y$ ,  $t(q, y, a, r^*) = \tilde{t}(q, y^*, a, r^*) = v(q, a) - s(q, a, r^*)$  where the first equality follows from equation 12 and the second equality follows from equation 11.

A similar argument can be made for residence status. In particular, suppose that for  $r^{**} \neq r^*$ :

$$\tilde{t}(q, y^*, a, r^{**}) + s(q, a, r^{**}) < \tilde{t}(q, y^*, a, r^*) + s(q, a, r^*)$$

Then colleges strictly prefer to admit students with residence status  $r^*$ , so if markets clear

it must be the case that students with residence status  $r^{**}$  do not want to attend. But then we can raise tuition for those students to the point at which college revenue (tuition plus subsidies) is equated across  $r^{**}$  and  $r^*$ .

Thus, we have established that a competitive equilibrium can be supported by a tuition schedule  $t(q, a, r)$  that is independent of income, and a revenue function  $v(q, a)$  that is independent of both income and residence status.

## Proof of Proposition 2

In this proof we show that for any equilibrium with a potentially nonlinear revenue schedule, we can construct an alternative linear revenue schedule that is consistent with the equilibrium allocation. Thus, any equilibrium can be supported by a revenue schedule that is linear with respect to ability. The proof is simple if there are no missing markets (*i.e.*, if  $\eta_a(q) > 0, \forall a$ ), but slightly more involved when only a subset of ability types attend colleges of a given quality, in which case there may be a set of tuition schedules consistent with the same allocation. We show that there must be a linear tuition schedule within this set.

*Proof.* Fix a college quality  $q$  such that there is positive supply,  $\chi(q) > 0$ . Denote the set of active college markets of quality  $q$  by  $A^+(q) = \{a : \eta_a(q) > 0\}$ . This set is nonempty given that  $\chi(q) > 0$ . Let  $a_{\max}$  and  $a_{\min}$  be the maximum and minimum elements in this set.

Case 1:  $A^+(q)$  is not a singleton set (so that  $a_{\max} > a_{\min}$ ):

In this case, define

$$d(q) = -\frac{v(q, a_{\max}) - v(q, a_{\min})}{a_{\max} - a_{\min}}$$

$$b(q) = v(q, a_{\min}) + d(q)(a_{\min} - a_1).$$

Now we claim that for any  $a_i$ ,

$$v(q, a_i) = b(q) - d(q)(a_i - a_1).$$

We prove this by contradiction. Suppose that for some  $j$

$$v(q, a_j) > b(q) - d(q)(a_j - a_1).$$

Then one can show that a college can increase profits by substituting in students of type  $a_j$  and substituting out a combination of students of types  $a_{\min}$  and  $a_{\max}$  to maintain the same average ability level in college, and can continue doing so until either  $\eta_{a_{\min}}$  or  $\eta_{a_{\max}}$  is zero. But this is a contradiction to the assumption that both  $a_{\max}$  and  $a_{\min}$  belong to the set of active markets  $A^+(q)$ . Thus, it must be that

$$v(q, a_j) \leq b(q) - d(q)(a_j - a_1) \text{ for any } j.$$

Now suppose that  $v(q, a_j) < b(q) - d(q)(a_j - a_1)$  for some  $j$ . It must then be that  $\eta_{a_j}(q) = 0$  in equilibrium. Otherwise, the college could shift admissions from  $a_j$  students to other ability levels, maintaining the desired average ability and making greater profit. Thus, in equilibrium it must be the case that both the supply and the demand in this particular market  $(q, a_j)$  are zero. Now replace the tuition value  $v(q, a_j)$  with  $\tilde{v}(q, a_j)$

defined by

$$\tilde{v}(q, a_j) = b(q) - d(q)(a_j - a_1).$$

Note that at the new level of tuition  $\tilde{v}(q, a_j)$ , college demand for students will still be zero (because colleges are indifferent between admitting students with ability  $a_j$  or a group of students with average ability  $a_j$ ). And the supply of students is zero as well because it is now more costly for the households to pick this college  $\tilde{v}(q, a_j) > v(q, a_j)$ . Thus the market still clears under the new revenue level  $\tilde{v}(q, a_j)$ . Thus, without loss of generality we can treat  $\tilde{v}(q, a_j)$  as the equilibrium revenue value. And we have finished the first part of the proof.

Case 2:  $A^+(q) = \{a : \eta_a(q) > 0\}$  is a singleton set:

In this case, first let  $a_m$  be the unique element of  $A^+(q)$ . Define a set of discount rates  $D^<(q) = \left\{d_i(q), i < m : d_i(q) = -\frac{v(q, a_i) - v(q, a_m)}{a_i - a_m}\right\}$ , which is the set of revenue slopes between  $a_m$  and other ability levels lower than  $a_m$ . If this set is nonempty, pick the greatest element in this set  $D^<(q)$ ,  $d_n(q)$ , and denote the associated  $a$  as  $a_n$ . Define

$$b_n(q) = v(q, a_m).$$

Now we claim that for any  $a_j$ , it must be the case that

$$v(q, a_j) \leq b_n(q) - d_n(q)(a_j - a_m).$$

To see this, note that for any  $a_j < a_m$ , by definition the slope  $d_j(q) \leq d_n(q)$ , and thus

$$\begin{aligned} v(q, a_j) &= v(q, a_m) - d_j(q)(a_j - a_m) \\ &= b_n(q) - d_j(q)(a_j - a_m) \\ &\geq b_n(q) - d_n(q)(a_j - a_m), \end{aligned}$$

where the last inequality holds because  $a_j - a_m < 0$ .

Next, for any  $a_j > a_m$ , we can show this by contradiction. Suppose that

$$v(q, a_j) > b_n(q) - d_n(q)(a_j - a_m).$$

Then the college can use a mix of  $a_j$  and  $a_n$  students to replicate ability level  $a_m$  (since  $a_j > a_m > a_n$ , such a mix is feasible). This yields greater profit for the college. But this is a contradiction to optimality. Thus, we proved that for any  $a_j$

$$v(q, a_j) \leq b_n(q) - d_n(q)(a_j - a_m).$$

And similarly to the first part of the proof, we can replace  $v(q, a_j)$  with

$$\tilde{v}(q, a_j) = b_n(q) - d_n(q)(a_j - a_m)$$

and maintain market clearing. The very last step is to show that even when the set  $D^<(q)$

is empty, we still have a linear tuition schedule. To see this, define

$$D^>(q) = \left\{ d_i(q), i > m : d_i(q) = -\frac{v(q, a_i) - v(q, a_m)}{a_i - a_m} \right\}.$$

This set must be nonempty given that  $D^<(q)$  is empty. Pick the smallest element in this set and denote it  $d_l(q)$  with associated ability level  $a_l(q)$ . Also define

$$b_l(q) = v(q, a_m).$$

Now we claim that for any  $a_j$ ,

$$v(q, a_j) \leq b_l(q) - d_l(q)(a_j - a_m)$$

The alternative case in which this inequality is not satisfied would violate the fact that  $d_l(q)$  is the smallest element of  $D^>(q)$  and that  $D^<(q)$  is empty. Thus, again similarly to the first part of the proof, we can replace  $v(q, a_j)$  with

$$\tilde{v}(q, a_j) = b_n(q) - d_n(q)(a_j - a_m)$$

and all the market-clearing conditions are satisfied. This completes the proof.  $\square$

For the proofs of Propositions 3 and 4, we consider a stripped-down version of the model. We shut down all government policies and transfers so that college revenue  $v(\cdot)$  is exactly equal to tuition  $t(\cdot)$ . Hence the tuition function only depends on  $q$  and  $a$ . We also abstract from dropout risk and income-based subsidies on the household side, so that the household's problem becomes:

$$\begin{aligned} & \max_{c \geq 0, q \in \mathcal{Q}} \mathbb{E} [u(c, q)] \\ & \text{s.t.} \\ & c + t(q, a) = y - \mathbb{1}_{\{q > 0\}} \omega. \end{aligned}$$

In this environment, a competitive equilibrium is defined by a family of functions  $\{\chi(q), c(y, a), q(y, a), \eta(q, a), e(q), t(q, a)\}$  such that households maximize, colleges maximize and earn zero profits, and all markets clear.

For the purpose of proving equilibrium existence, we define a notion of **quasi-equilibrium**. In a quasi-equilibrium, the households' decision rules  $\{c(y, a), q(y, a)\}$  satisfy quasi-optimization: for any alternative allocations  $c'(y, a), q'(y, a)$  such that  $\mathbb{E} u(c'(y, a), q'(y, a)) > \mathbb{E} u(c(y, a), q(y, a))$ , it must be the case that  $c'(y, a) + t(q'(y, a), a) \geq y - \mathbb{1}_{\{q > 0\}} \omega$ .

An equilibrium is necessarily a quasi-equilibrium but not vice versa.

### Proof of Proposition 3

This proof consists of two steps. In the first step, we establish that a competitive equilibrium exists when there are finitely many types of colleges, i.e. when the set  $\Omega = \{q_1, q_2, \dots, q_n\}_{n < \infty}$  contains only a finite number of discrete elements. This part of the proof draws on Ellickson et al. (1999), which establishes equilibrium existence in a general club model setting. However, there are two subtle differences between our model and that of Ellickson et al. (1999). First, Ellickson et al. (1999) only allow for a finite number of types of clubs, while we are ultimately interested in a situation where colleges can enter at any quality level desired, and the equilibrium quality distribution is potentially continuous. Second, Ellickson et al. (1999) assume multiple private consumption goods, while we only have one type of consumption good.

#### Step 1: Prove existence with finite types of clubs

We start by verifying that Theorem 6.1 of Ellickson et al. (1999) holds in our environment when  $\Omega$  is a finite discrete set:

*If the agents' endowments are desirable and uniformly bounded from above, then a quasi-equilibrium exists. (Theorem 6.1, page 1201)*

Uniform boundedness of endowments is satisfied, since the income distribution is defined over a compact set  $[y_{\min}, y_{\max}]$ . Next we check the desirability of endowment

requirement.

The endowments are desirable if for every household, consuming his endowment and no club membership gives strictly higher utility than consuming no private consumption and any membership:

$$\mathbb{E}[u(y, 0)] > \mathbb{E}[u(0, q)], \forall y \in [y_{\min}, y_{\max}] > 0, q \in \Omega$$

This condition is satisfied in our setting because of the reservation utility of no college  $\kappa > 0$ , while consuming no private consumption goods give negative infinite utility.

We then need to show that a quasi-equilibrium is indeed an equilibrium. In Ellickson et al. (1999) this is done through a "club irreducibility condition", which depends on there being more than one type of private consumption good. Given that our model only has one type of consumption good, we cannot use their condition to verify that an equilibrium exists. Instead we exploit the local non-satiation property of our utility function to directly establish existence. To see this, suppose that the quasi-equilibrium is not an equilibrium. Then it must be the case that there exists an alternative consumption bundle  $c'(y, a), q'(y, a)$  such that  $\mathbb{E}u(c'(y, a), q'(y, a)) > \mathbb{E}u(c(y, a), q(y, a))$  and

$$c'(y, a) + t(q'(y, a), a) = y - \mathbb{1}_{\{q'(y, a) > 0\}}\omega.$$

Suppose that  $c'(y, a) > 0$  (to be verified later). By continuity we can reduce  $c'(y, a)$  a little bit and the resulting allocation (call it)  $(c'', q')$  lies strictly inside the budget constraint:

$$c''(y, a) + t(q'(y, a), a) < y - \mathbb{1}_{\{q'(y, a) > 0\}}\omega$$

And is strictly preferred to the competitive allocation:  $\mathbb{E}u(c''(y, a), q'(y, a)) > \mathbb{E}u(c(y, a), q(y, a))$ . This contradicts the quasi-optimization condition that the competitive allocation is strictly preferred to any allocations strictly inside the budget constraint.

The last thing we need to verify is that  $c'(y, a)$  is indeed strictly positive. This can be guaranteed with the assumption that the household income distribution has a strictly positive support  $y_{\min} > 0$ . Suppose that  $c'(y, a) = 0$ . Then by the form of utility function we know that  $\mathbb{E}u(c'(y, a), q'(y, a)) = -\infty$ . Now given that  $y \geq y_{\min} > 0$ , the consumption bundle  $(y, 0)$  is feasible. Therefore  $\mathbb{E}u(c(y, a), q(y, a)) \geq \mathbb{E}u(y, 0) > -\infty$ . This contradicts the assumption that  $\mathbb{E}u(c'(y, a), q'(y, a)) > \mathbb{E}u(c(y, a), q(y, a))$ . Thus  $c'(y, a)$  must be strictly positive. This concludes the existence proof when the set  $\Omega$  has finitely many elements.

## Step 2: Extending the existence result to a continuum of college types

Next, we extend existence to a continuum of college types, i.e.,  $\Omega = [0, q_{\max}]$ . Note that Caucutt (2001) shows that an equilibrium exists by introducing lotteries to convexify individual consumption sets. However, the step in her proof going from finitely many college types to infinitely many college types does not depend on the assumption of convex consumption sets and can therefore be adapted here.

We first show that an equilibrium exists if the aggregate technology set is restricted so that only a finite number of school types are permitted to operate. There the first step of the proof applies and we can show that an equilibrium exists. These economies are referred to as  $r^{\text{th}}$  approximate economies, where  $r$  denotes elements in this restricted set  $\Omega_r$ . We then show a convergent subsequence of these economies exists and that the limit is an equilibrium for an economy with a continuum of college types.

More precisely, we pick the  $r^{\text{th}}$  approximate economy as follows. We set the college quality set  $\Omega_r \subset \Omega$  such that all points in  $\Omega$  lies in a  $\frac{1}{r}$ -neighborhood of at least one point in  $\Omega_r$ . For the  $r^{\text{th}}$  approximate economy, we know that an equilibrium exists. Denote it by  $t^r(q, a)$  and  $x^r$ , where  $x^r$  is abbreviated notation for the allocation associated with the competitive equilibrium  $\{\chi(q), c(y, a), q(y, a), \eta(q, a), e(q)\}_r$ . Then our goal is to show that the sequence  $\{t^r, x^r\}$  converges.

### Step 2.1: Uniform Convergence of the Tuition Function

We first show that the tuition function converges uniformly. This involves showing that the tuition functions are bounded uniformly and are Lipschitz continuous.

To show that the tuition functions are bounded uniformly, it is useful to note from Proposition 2 that the tuition functions take the following form:

$$t^r(q, a) = b^r(q) - d^r(q)(a - a_{\min})$$

Given that the ability space is bounded, it suffices to show that the functions  $b^r$  and  $d^r$  are bounded uniformly. We first show that  $d^r(q)$  is bounded. Note from the college's FOC (equation 8):

$$d^r(q) = \frac{\theta}{1 - \theta} \frac{e(q)}{\bar{a}(q)}$$

Given that the  $(a, q)$  space is bounded with  $a_{\min} > 0$ , the expenditure  $e(q)$  is also implicitly bounded according to the production function  $q = \bar{a}^\theta e^{1-\theta}$ . This shows that  $d^r(q)$  is bounded.

Now we show that  $b^r(q)$  is also bounded uniformly. We divide it into two cases.



Suppose that  $\chi(q) > 0$ . Then zero profit condition must hold:

$$b^r(q) - d^r(q)(\bar{a} - a_{\min}) - e - \phi = 0$$

$$b^r(q) = d^r(q)(\bar{a} - a_{\min}) + e + \phi$$

Since all objects on the right-hand-side are uniform bounded,  $b^r(q)$  is uniform bounded.

Next we check the case  $\chi(q) = 0$ . In this case the college makes non-positive profit:

$$b^r(q) - d^r(q)(\bar{a} - a_{\min}) - e - \phi \leq 0$$

This provides an upper bound on  $b^r(q)$ . To establish a lower bound of  $b^r(q)$ , we focus on the  $(a_{\min}, y_{\min})$  household. We impose that the opportunity cost of going to college  $\omega$  is larger than  $y_{\min}$ , which implies that the household at  $y_{\min}$  will optimally choose to consume his endowment  $c = y_{\min}$ . Otherwise his utility would be negative infinity as he would consume negative consumption. Thus

$$\log(y_{\min}) + \phi \log(\kappa) \geq \log(y_{\min} - b^r(q)) + \phi \log(\kappa + q)$$

$$y_{\min} \left( \frac{\kappa}{\kappa + q} \right)^\phi \geq y_{\min} - b^r(q)$$

$$b^r(q) \geq y_{\min} \left( 1 - \left( \frac{\kappa}{\kappa + q} \right)^\phi \right) \geq 0 \text{ as } q \geq 0$$

Thus we have found a uniform lower bound for  $b^r(q)$ . This concludes the proof that the tuition function is uniformly bounded.

We next show that the tuition function  $t^r(q, a)$  is Lipschitz continuous. That is, there exists a  $k$  independent of  $r$  such that for any  $q, q', a, a'$

$$|t^r(q, a) - t^r(q', a')| \leq k |(q, a) - (q', a')|$$

where the metric  $|(q, a) - (q', a')|$  is given by  $|q - q'| + |a - a'|$ . By triangular inequality, it suffices to show that the tuition function is Lipschitz continuous with respect to  $q$  and  $a$  respectively:

$$\begin{aligned} |t^r(q, a) - t^r(q', a')| &= |t^r(q, a) - t^r(q, a') + t^r(q, a') - t^r(q', a')| \\ &\leq |t^r(q, a) - t^r(q, a')| + |t^r(q, a') - t^r(q', a')| \end{aligned}$$

For the first part, we have by the linearity property of the tuition schedule:

$$|t^r(q, a) - t^r(q, a')| = |d^r(q)| |a - a'|$$

By the uniform boundedness of  $d^r(q)$ , we know that  $t^r$  is Lipschitz continuous with respect to  $a$ .

We next show that the tuition function is Lipschitz continuous with respect to  $q$ . We invoke the household's indifference condition between  $q$  and  $q'$ . There must exist a level of income  $\bar{y}$  such that:

$$\log(\bar{y} - t^r(q, a)) + \varphi \log(\kappa + q) = \log(\bar{y} - t^r(q', a)) + \varphi \log(\kappa + q')$$

$$\bar{y}(\kappa + q)^\varphi - t^r(q, a)(\kappa + q)^\varphi = \bar{y}(\kappa + q')^\varphi - t^r(q', a)(\kappa + q')^\varphi$$

$$t^r(q', a)(\kappa + q')^\varphi - t^r(q', a)(\kappa + q)^\varphi + t^r(q', a)(\kappa + q)^\varphi - t^r(q, a)(\kappa + q)^\varphi = \bar{y} \left( (\kappa + q')^\varphi - (\kappa + q)^\varphi \right)$$

Thus

$$[t^r(q, a) - t^r(q', a)](\kappa + q)^\varphi = (\bar{y} - t^r(q', a)) \left( (\kappa + q')^\varphi - (\kappa + q)^\varphi \right)$$

Or

$$|t^r(q, a) - t^r(q', a)| = |(\bar{y} - t^r(q', a))| \left| \frac{(\kappa + q')^\varphi - (\kappa + q)^\varphi}{(\kappa + q)^\varphi} \right|$$

Given that income  $\bar{y}$  is drawn from a compact set  $[y_{\min}, y_{\max}]$  and  $t^r(q', a)$  is bounded uniformly in  $r$ , we have

$$|t^r(q, a) - t^r(q', a)| \leq K \left| \frac{(\kappa + q')^\varphi}{(\kappa + q)^\varphi} - 1 \right|$$

for some constant  $K$ . Now given that the function  $\frac{(\kappa + q')^\varphi}{(\kappa + q)^\varphi} - 1$  is continuously differentiable and hence Lipschitz continuous. This concludes the proof that tuition function  $t^r(q, a)$  is Lipschitz continuous with respect to  $q$ .

The uniform boundedness and Lipschitz continuity guarantees that the family of tuition functions  $\{t^r(q, a)\}_r$  converges uniformly to some tuition function  $\{t^*(q, a)\}$  as  $r \rightarrow \infty$ .

### Step 2.2: Convergence of allocation

We then show that the allocation converges. Note that the consumption sets for each household are closed and bounded (note that convexity is not required here). Hence the aggregate consumption set is closed and bounded. The aggregate production set is also closed and bounded, as quality and ability are closed and bounded sets. Therefore there

exists a subsequence of  $\{x^r\}_r$  that converges to some limit  $\{x^*\}$  as  $r \rightarrow \infty$ .

**Step 2.3: Verify that the limit equilibrium satisfies equilibrium conditions**

Having established the existence of the limit equilibrium  $\{t^*, x^*\}$ , as a last step we need to verify that the limit equilibrium satisfies the equilibrium conditions. We first verify that it satisfies household optimality. Because  $x^r$  is an allocation associated with  $t^r$ , it must satisfy budget feasibility:

$$c^r(y, a) + t^r(q^r(y, a), a) = y - \mathbb{1}_{\{q^r(y, a) > 0\}} \omega$$

Taking the limit we have that  $x^*$  is feasible under  $t^*$  :

$$c^*(y, a) + t^*(q^*(y, a), a) = y - \mathbb{1}_{\{q^*(y, a) > 0\}} \omega$$

$c^*(y, a), q^*(y, a)$  are also optimal given  $t^*(q, a)$ . Otherwise the household optimality condition would be violated for some  $r$  sufficiently large.

We next verify college optimality. This can be done easily by noticing that colleges' profit function is continuous with respect to the tuition function and choices. Likewise we can also verify the zero profit condition and market clearing conditions.

### Proof of Proposition 4

We first define a feasible allocation in the benchmark economy.

**Definition.** A feasible allocation is a set of functions  $\{\chi(q), c(y, a), q(y, a), \eta(q, a), e(q)\}$  such that:

1. The allocation of final output is feasible:

$$\sum_a \mu_a \int_0^\infty c(y, a) dF_a(y) + \int_0^\infty e(q) d\chi(q) + (1 - \chi(0))(\omega + \phi) = \sum_a \mu_a \int_0^\infty y dF_a(y).$$

2. The allocation of students is feasible. For all  $a$  and  $Q \subset \mathbb{R}^+$ ,

$$\mu_a \int_0^\infty \mathbb{1}_{\{q(y, a) \in Q\}} dF_a(y) = \int_Q \eta(q, a) d\chi(q),$$

where  $\mathbb{1}_{\{\cdot\}}$  is an indicator function.

3. The allocation of quality is feasible. For all  $Q$ ,

$$\int_Q \sum_a \mu_a \int_0^\infty \mathbb{1}_{\{q(y, a) \in Q\}} q(y, a) dF_a(y) d\chi(q) = \int_Q \left( \sum_a \eta(q, a) a \right)^\theta e(q)^{1-\theta} d\chi(q)$$

**Definition.** An allocation  $\{\chi(q), c(y, a), q(y, a), \eta(q, a), e(q)\}$  is Pareto optimal if 1) it is a feasible allocation and 2) there does not exist a feasible allocation  $\{\chi'(q), c'(y, a), q'(y, a), \eta'(q, a), e'(q)\}$  such that  $u(c'(y, a), q'(y, a)) \geq u(c(y, a), q(y, a))$  for almost every  $(y, a)$  and  $u(c'(y, a), q'(y, a)) > u(c(y, a), q(y, a))$  for some set  $(y, a)$  of positive measure.

We now prove the First Welfare Theorem. The proof here closely mirrors the standard proof of the Welfare Theorem.

**Theorem.** Assume that  $u$  exhibits local nonsatiation. If  $\{\chi(q), c(y, a), q(y, a), \eta(q, a), e(q)\}$  is a competitive equilibrium allocation, then it is Pareto efficient.

**Proof.** Suppose to the contrary that there exists an alternative feasible allocation  $\{\chi'(q), c'(y, a), q'(y, a), \eta'(q, a), e'(q)\}$  such that  $u(c'(y, a), q'(y, a)) \geq u(c(y, a), q(y, a))$  for almost every  $(y, a)$  and  $u(c'(y, a), q'(y, a)) > u(c(y, a), q(y, a))$  for some positive measure set of  $(y, a)$ . Denote  $t(q, a)$  the equilibrium tuition function associated with the competitive equilibrium allocation. Then, from the local non-satiation assumption, we know that the alternative allocation must lie outside households' budget set:

$$c'(y, a) + t(q'(y, a), a) \geq y - \mathbb{1}_{\{q'(y, a) > 0\}} \omega \text{ for almost every } (y, a).$$

Otherwise,  $c(y, a), q(y, a)$  would not be individually rational given the tuition functions. In addition,

$$c'(y, a) + t(q'(y, a), a) > y - \mathbb{1}_{\{q'(y, a) > 0\}} \omega \text{ for some positive measure set.}$$

Summing up the above equations across households of different abilities and income, we get

$$\sum_a \mu_a \int c'(y, a) dF_a(y) + \sum_a \mu_a \int t(q'(y, a), a) dF_a(y) > \sum_a \mu_a \int y dF_a(y) - (1 - \chi'(0))\omega. \quad (13)$$

Note that under the alternative feasible allocation, aggregate enrollment is given by  $1 - \chi'(0)$ .

Now turn to the college sector. Since the equilibrium allocation maximizes the colleges' profit under the competitive price vector, the alternative allocation must be (weakly) inferior to the competitive equilibrium allocation under the competitive tuition schedule. Thus, the colleges must make nonpositive profit:

$$\sum_a \eta'(q, a)t(q, a) - e'(q) - \phi \leq 0.$$

Since this equation holds for all colleges, the aggregate profit made by the college sector must be nonpositive, which in turn implies that the aggregate tuition revenue is no greater than the college expenditure:

$$\sum_a \int t(q'(y, a), a) dF_a(y) \leq \int_0^{q_{\max}} e'(q) d\chi'(q) + (1 - \chi'(0))\phi. \quad (14)$$

But eqs. 13 and 14 together imply that the final goods resource constraint is violated under the alternative allocation:

$$\begin{aligned} & \sum_a \mu_a \int c'(y, a) dF_a(y) + \int_0^{q_{\max}} e'(q) d\chi'(q) + (1 - \chi'(0))(\omega + \phi) \\ & \geq \sum_a \mu_a \int c'(y, a) dF_a(y) + \sum_a \int t(q'(y, a), a) dF_a(y) + (1 - \chi'(0))\omega \\ & > \sum_a \mu_a \int y dF_a(y), \end{aligned}$$

where the first weak inequality follows from equation 14 and the second strict inequality follows from equation 13. Thus we have a contradiction. This establishes the Pareto-

optimality of the competitive allocation.

### Proof of Proposition 5

**Theorem.** Suppose education is a pure club good ( $\theta = 1$ ) and there are only two ability levels  $a_h$  and  $a_l$  and household utility is given by

$$\log(c) + \log(\kappa + q),$$

and the income distribution for both high and low ability is uniform in some interval  $[\mu_y - \frac{1}{2}\Delta_y, \mu_y + \frac{1}{2}\Delta_y]$ . Then the college distribution is given by

$$\begin{aligned}\chi(Q) &= \frac{2}{a_h - a_l} \frac{2}{4 + \pi} \int_Q \left[ (1 - \eta(q))^2 + \eta(q)^2 \right]^{-2} dq \quad \forall Q \subset (a_l, a_h) \\ \chi(a_h) &= \chi(a_l) = \frac{2}{4 + \pi},\end{aligned}$$

and the tuition function is given by

$$t(q, a) = \mu_y \frac{q - a}{\kappa + q} \left[ 1 - \frac{2}{4 + \pi} \frac{\Delta_y}{\mu_y} \arctan(1 - 2\eta(q)) \right], i = h, l.$$

**Proof.** As a first step, the household problem given income  $y$  is

$$\max_{c, q} \log(c) + \log(\kappa + q)$$

$$c + t(q, a) \leq y.$$

Assuming that  $t(q, a)$  is differentiable with respect to  $q$  (verified later), the first-order condition with respect to  $q$  for households of ability  $a$  is

$$t'(q, a) = \frac{y - t(q, a)}{\kappa + q}.$$

Denoting by  $y(q)$  the income of the households attending colleges of quality  $q$  in equilibrium, we have

$$\begin{aligned}t'(q, a) &= \frac{y(q) - t(q, a)}{\kappa + q} \\ &= -\frac{1}{\kappa + q} t(q, a) + \frac{y(q)}{\kappa + q}.\end{aligned}$$

This is a linear ordinary differential equation that can be solved using the integrating

factor method. Define the integrating factor for low ability tuition  $v^l(q)$  as

$$\begin{aligned} v^l(q) &= \int_{a_l}^q \frac{1}{\kappa + q'} dq' \\ &= \log \frac{\kappa + q}{\kappa + a_l}. \end{aligned}$$

Thus,

$$\exp(v^l(q)) = \frac{\kappa + q}{\kappa + a_l}$$

$$\begin{aligned} \exp(v^l(q)) t'(q, a) + \frac{1}{\kappa + q} \exp(v^l(q)) t(q, a_l) &= \exp(v^l(q)) \frac{y(q)}{\kappa + q} \\ \left[ \exp(v^l(q)) t(q, a_l) \right]' &= \exp(v^l(q)) \frac{y(q)}{\kappa + q} \\ \int_{a_l}^q \left[ \exp(v^l(q)) t(q, a_l) \right]' dq &= \int_{a_l}^q \exp(v^l(q)) \frac{y(q)}{\kappa + q} dq \\ \exp(v^l(q)) t(q, a_l) - \exp(a_l) t(q, a_l) &= \int_{a_l}^q \exp(v^l(q)) \frac{y(q)}{\kappa + q} dq \\ &= \int_{a_l}^q \frac{y(q)}{\kappa + a_l} dq. \end{aligned}$$

We know from the zero profit condition that

$$t(a_l, a_l) = 0$$

which means that a college of quality  $a_l$  (which must consists of  $a_l$  students only) must charge those  $a_l$  kids zero tuition. Thus

$$\begin{aligned} \exp(v^l(q)) t(q, a_l) &= \int_{a_l}^q \frac{y(q')}{\kappa + a_l} dq' \\ t(q, a_l) &= \exp(-v^l(q)) \int_{a_l}^q \frac{y(q')}{\kappa + a_l} dq' \\ &= \frac{\kappa + a_l}{\kappa + q} \int_{a_l}^q \frac{y(q')}{\kappa + a_l} dq' \\ &= \int_{a_l}^q \frac{y^l(q')}{\kappa + q} dq'. \end{aligned}$$



Likewise, the integrating factor for the high-ability type is given by

$$\begin{aligned} v^h(q) &= \int_{a_h}^q \frac{1}{\kappa + q'} dq' \\ &= \log \frac{\kappa + q}{\kappa + a_h}. \end{aligned}$$

And we can follow the same procedure and obtain an expression for the high-ability tuition function:

$$\exp(v^h(q)) t(q, a_h) - \exp(v^h(a_h)) t(a_h, a_h) = \int_{a_h}^q \exp(v^h(q)) \frac{y(q)}{\kappa + q} dq.$$

The zero profit condition for the  $q = a_h$  college implies

$$t(a_h, a_h) = 0.$$

Thus,

$$\begin{aligned} t(q, a_h) &= \int_{a_h}^q \frac{y(q)}{\kappa + q} dq \\ &= - \int_q^{a_h} \frac{y^h(q)}{\kappa + q} dq. \end{aligned}$$

Now we derive the income function  $y^a(q)$ , given uniformly distributed income and any college distribution function  $\chi(q)$  :

$$\begin{aligned} y^h(q) &= \mu_y + \frac{1}{2}\Delta_y - \Delta_y \int_q^{a_h} \chi(q') \frac{q' - a_l}{a_h - a_l} dq' \\ y^l(q) &= \mu_y + \frac{1}{2}\Delta_y - \Delta_y \int_q^{a_h} \chi(q') \frac{a_h - q'}{a_h - a_l} dq'. \end{aligned}$$

Now, we would like to solve for  $\chi(q)$  from the zero profit condition  $\pi(q) = 0$ . We conjecture that there is a strictly positive measure of high-ability students going to colleges of quality  $q = a_h$ . Denote that mass  $\chi(a_h)$ . Thus,

$$y^h(q) = \mu_y + \frac{1}{2}\Delta_y - \Delta_y \left( \int_q^{a_h} \chi(q') \frac{q' - a_l}{a_h - a_l} dq' + \chi(a_h) \right).$$

Write out the expression for  $\pi(q)$  :

$$\begin{aligned}
\pi(q) &= \frac{q - a_l}{a_h - a_l} t^h(q) + \frac{a_h - q}{a_h - a_l} t^l(q) \\
&= -\frac{q - a_l}{a_h - a_l} \int_q^{a_h} \frac{y^h(q')}{\kappa + q} dq' + \frac{a_h - q}{a_h - a_l} \int_{a_l}^q \frac{y^l(q')}{\kappa + q} dq' \\
&= 0.
\end{aligned}$$

Canceling out some terms, we have that for any  $q$

$$(a_h - q) \int_{a_l}^q y^l(q') dq' - (q - a_l) \int_q^{a_h} y^h(q') dq' = 0$$

Substitute in expressions for  $y^l(q)$  and  $y^h(q)$ :

$$\begin{aligned}
&(a_h - q) \left( \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_{q'}^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) dq' \right) - \\
&(q - a_l) \left( \int_q^{a_h} \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \left( \int_{q'}^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) dq' \right) = 0.
\end{aligned}$$

Differentiate with respect to  $q$  :

$$\begin{aligned}
&- \left( \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_x^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) dq' \right) \\
&+ (a_h - q) \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_q^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) \\
&- \left( \int_q^{a_h} \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \left( \int_x^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) dq \right) \\
&+ (q - a_l) \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \left( \int_q^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) = 0.
\end{aligned}$$

Differentiate again with respect to  $q$  :

$$\begin{aligned}
& - \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_q^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) \\
& - \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_q^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) \\
& + (a_h - q) \Delta_y \chi(q) \frac{a_h - q}{a_h - a_l} + \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \left( \int_x^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) \\
& + \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \left( \int_q^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) + (q - a_l) \Delta_y \left( \chi(q) \frac{q - a_l}{a_h - a_l} \right) = 0.
\end{aligned}$$

Collect terms:

$$\begin{aligned}
& -2 \left( -\Delta_y \int_q^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) + (a_h - q) \Delta_y \chi(q) \frac{a_h - q}{a_h - a_l} \\
& 2 \left( -\Delta_y \left( \int_q^{a_h} \chi(x) \frac{x - a_l}{a_h - a_l} dx + \chi(a_h) \right) \right) + (q - a_l) \Delta_y \left( \chi(q) \frac{q - a_l}{a_h - a_l} \right) = 0.
\end{aligned}$$

Note that  $\Delta_y$  can be factored out, and we arrive at a functional equation  $\chi(q)$  that is independent of the income distribution parameters:

$$\begin{aligned}
& 2 \int_q^{a_h} \chi(x) (a_h - x) dx + (a_h - q)^2 \chi(q) \\
& -2 \left( \int_q^{a_h} \chi(x) (x - a_l) dx + \chi(a_h) [a_h - a_l] \right) + (q - a_l)^2 \chi(q) \\
& = 0
\end{aligned}$$

Thus, we have the following integral equation:

$$\begin{aligned}
& \left[ (a_h - q)^2 + (q - a_l)^2 \right] \chi(q) + 2 \int_q^{a_h} \chi(x) (a_h + a_l - 2x) dx = 2\chi(a_h) [a_h - a_l] \\
& \chi(q) = \frac{-2}{\left[ (a_h - q)^2 + (q - a_l)^2 \right]} \int_{a_h}^q \chi(x) (2x - a_h - a_l) dx + \frac{2\chi(a_h) [a_h - a_l]}{(a_h - q)^2 + (q - a_l)^2}.
\end{aligned}$$

This is a **Volterra equation of the second type with degenerate kernels**, which happens

to have an analytical solution. Define the following objects:

$$\begin{aligned} g(q) &= \frac{-2}{\left[(a_h - q)^2 + (q - a_l)^2\right]} \\ h(x) &= (2x - a_h - a_l) \\ f(q) &= \frac{2\chi(a_h)[a_h - a_l]}{(a_h - q)^2 + (q - a_l)^2}. \end{aligned}$$

and we have

$$\chi(q) = f(q) + \int_{a_h}^q R(q, x) f(x) dx$$

where

$$\begin{aligned} R(q, x) &= g(q) h(x) \exp \left[ \int_x^q g(s) h(s) ds \right] \\ &= \frac{-2(2x - a_h - a_l)}{\left[(a_h - q)^2 + (q - a_l)^2\right]} \exp \left[ \int_x^q \frac{-2(2s - a_h - a_l)}{\left[(a_h - s)^2 + (s - a_l)^2\right]} ds \right] \\ &= \frac{-2(2x - a_h - a_l)}{\left[(a_h - q)^2 + (q - a_l)^2\right]} \exp \left[ - \int_x^q \frac{1}{\left[(a_h - s)^2 + (s - a_l)^2\right]} d \left[ (a_h - s)^2 + (s - a_l)^2 \right] \right] \\ &= \frac{-2(2x - a_h - a_l)}{\left[(a_h - q)^2 + (q - a_l)^2\right]} \exp \left[ - \log \frac{(a_h - q)^2 + (q - a_l)^2}{(a_h - x)^2 + (x - a_l)^2} \right] \\ &= \frac{-2(2x - a_h - a_l) \left( (a_h - x)^2 + (x - a_l)^2 \right)}{\left[ (a_h - q)^2 + (q - a_l)^2 \right]^2}. \end{aligned}$$

Now,

$$\begin{aligned} &\int_{a_h}^q R(q, x) f(x) dx \\ &= \int_{a_h}^q \frac{-2(2x - a_h - a_l) \left( (a_h - x)^2 + (x - a_l)^2 \right)}{\left[ (a_h - q)^2 + (q - a_l)^2 \right]^2} \frac{2\chi(a_h)[a_h - a_l]}{(a_h - x)^2 + (x - a_l)^2} dx \\ &= \frac{-2\chi(a_h)[a_h - a_l]}{\left[ (a_h - q)^2 + (q - a_l)^2 \right]^2} \left[ (a_h - q)^2 + (q - a_l)^2 \right] \Big|_{a_h}^q \\ &= \frac{-2\chi(a_h)[a_h - a_l]}{\left[ (a_h - q)^2 + (q - a_l)^2 \right]^2} \left[ (a_h - q)^2 + (q - a_l)^2 - (a_h - a_l)^2 \right]. \end{aligned}$$

Thus,

$$\begin{aligned}
\chi(q) &= \frac{2\chi(a_h)[a_h - a_l]}{(a_h - q)^2 + (q - a_l)^2} + \frac{-2\chi(a_h)[a_h - a_l]}{\left[(a_h - q)^2 + (q - a_l)^2\right]^2} \left[(a_h - q)^2 + (q - a_l)^2 - (a_h - a_l)^2\right] \\
&= \frac{2\chi(a_h)[a_h - a_l]^3}{\left[(a_h - q)^2 + (q - a_l)^2\right]^2} \\
&= M(q)\chi(a_h),
\end{aligned}$$

where

$$M(q) = \frac{2[a_h - a_l]^3}{\left[(a_h - q)^2 + (q - a_l)^2\right]^2}.$$

Thus, we can derive the value of  $\chi(a_h)$  from

$$\begin{aligned}
\int_{a_l}^{a_h} \chi(q) \frac{q - a_l}{a_h - a_l} dq + \chi(a_h) &= 1 \\
\chi(a_h) &= \frac{1}{1 + \int_{a_l}^{a_h} M(q) \frac{q - a_l}{a_h - a_l} dq}.
\end{aligned}$$

Now,

$$\begin{aligned}
&\int_{a_l}^{a_h} M(q) \frac{q - a_l}{a_h - a_l} dq \\
&= 2[a_h - a_l]^2 \int_{a_l}^{a_h} \frac{(q - a_l)}{\left[(a_h - q)^2 + (q - a_l)^2\right]^2} dq \\
&= 2[a_h - a_l]^2 \frac{\frac{(a_h - a_l)(q - a_h)}{(a_h - q)^2 + (q - a_l)^2} + \arctan \frac{2q - a_l - a_h}{a_h - a_l}}{2(a_h - a_l)^2} \Big|_{a_l}^{a_h} \\
&= \frac{(a_h - a_l)(q - a_h)}{(a_h - q)^2 + (q - a_l)^2} + \arctan \frac{2q - a_l - a_h}{a_h - a_l} \Big|_{a_l}^{a_h} \\
&= 0 + \arctan(1) + 1 - \arctan(-1) \\
&= 1 + 2\arctan(1) \\
&= 1 + \frac{\pi}{2}.
\end{aligned}$$

Thus,

$$\chi(a_h) = \frac{1}{2 + \frac{\pi}{2}}$$

$$\chi(q) = \frac{2[a_h - a_l]^3}{\left[(a_h - q)^2 + (q - a_l)^2\right]^2} \frac{1}{2 + \frac{\pi}{2}}.$$

Now we know that all low-ability types not going to  $q > a_l$  will go to  $q = a_l$  college. That is given by

$$\begin{aligned} \chi(a_l) &= 1 - \int_{a_l}^{a_h} \chi(q') \frac{a_h - q'}{a_h - a_l} dq' \\ &= 1 - \chi(a_h) \int_{a_l}^{a_h} \frac{2[a_h - a_l]^2 (a_h - q')}{\left[(a_h - q')^2 + (q' - a_l)^2\right]^2} dq' \\ &= 1 - \chi(a_h) \left[ -\arctan \frac{2x - a_l - a_h}{a_l - a_h} - \frac{(a_l - a_h)(x - a_l)}{(a_h - x)^2 + (x - a_l)^2} \right] \Big|_{a_l}^{a_h} \\ &= 1 - \chi(a_h) (2 \arctan(1) + 1) \\ &= 1 - \frac{1}{2 + \frac{\pi}{2}} \left( \frac{\pi}{2} + 1 \right) \\ &= \frac{1}{2 + \frac{\pi}{2}}. \end{aligned}$$

Now we would like to derive closed forms for tuition functions:

$$\begin{aligned}
t(q, a_l) &= \int_{a_l}^q \frac{y^l(q')}{\kappa + q} dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y - \Delta_y \int_{q'}^{a_h} \chi(x) \frac{a_h - x}{a_h - a_l} dx \right) dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y - \frac{\Delta_y}{2 + \frac{\pi}{2}} \int_{q'}^{a_h} \frac{2[a_h - a_l]^2 (a_h - x)}{[(a_h - x)^2 + (x - a_l)^2]^2} dx \right) dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y - \frac{\Delta_y}{2 + \frac{\pi}{2}} \left( -\arctan \frac{2x - a_l - a_h}{a_l - a_h} - \frac{(a_l - a_h)(x - l)}{(a_h - x)^2 + (x - a_l)^2} \right) \Big|_{q'}^{a_h} \right) dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y + \frac{\Delta_y}{2 + \frac{\pi}{2}} \left( \arctan \frac{2a_h - a_l - a_h}{a_l - a_h} + \frac{(a_l - a_h)(a_h - a_l)}{(a_h - a_h)^2 + (a_h - a_l)^2} - \arctan \frac{2q' - a_l - a_h}{a_l - a_h} - \frac{(a_l - a_h)(q' - l)}{(a_h - x)^2 + (x - a_l)^2} \right) \right) dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{1}{2} \Delta_y + \frac{\Delta_y}{2 + \frac{\pi}{2}} \left( -\frac{\pi}{4} - 1 - \arctan \frac{2q' - a_l - a_h}{a_l - a_h} - \frac{(a_l - a_h)(q' - l)}{(a_h - q')^2 + (q' - a_l)^2} \right) \right) dq' \\
&= \frac{1}{\kappa + q} \int_{a_l}^q \left( \mu_y + \frac{\Delta_y}{2 + \frac{\pi}{2}} \left( -\arctan \frac{2q' - a_l - a_h}{a_l - a_h} - \frac{(a_l - a_h)(q' - l)}{(a_h - q')^2 + (q' - a_l)^2} \right) \right) dq' \\
&= \frac{1}{\kappa + q} \left[ \int_{a_l}^q adq' - \frac{\Delta_y}{2 + \frac{\pi}{2}} \int_{a_l}^q \arctan \frac{2x - a_l - a_h}{a_l - a_h} + \frac{(a_l - a_h)(x - l)}{(a_h - x)^2 + (x - a_l)^2} dx \right] \\
&= \frac{1}{\kappa + q} \left[ \mu_y (q - a_l) - \frac{\Delta_y}{2 + \frac{\pi}{2}} (x - a_l) \arctan \frac{2x - a_l - a_h}{a_l - a_h} \Big|_{a_l}^q \right] \\
&= \frac{(q - a_l)}{\kappa + q} \left[ \mu_y - \frac{\Delta_y}{2 + \frac{\pi}{2}} \arctan \frac{2q - a_l - a_h}{a_l - a_h} \right].
\end{aligned}$$

Following similar steps:

$$t(q, a_h) = -\frac{(a_h - q)}{\kappa + q} \left[ \mu_y - \frac{\Delta_y}{2 + \frac{\pi}{2}} \arctan \frac{2q - a_l - a_h}{a_l - a_h} \right].$$

Rearranging and plugging in the expressions for  $\eta(q)$ , we obtain the theorem.

### Proof of Proposition 6

This proposition establishes efficiency of equilibrium when education is treated as an investment good and there is a perfect credit market. Suppose, to the contrary, that the equilibrium allocation  $\{c_1(a, y), c_2(a, y), q(a, y), \chi(q), \eta_a(q), e(q)\}$  is not efficient. Then there exists an alternative allocation

$\{c'_1(a, y), c'_2(a, y), q'(a, y), \chi'(q), \eta'_a(q), e'(q)\}$  that is feasible and that Pareto dominates the equilibrium allocation. Then the household decision rules  $\{c'_1(a, y), c'_2(a, y), q'(a, y)\}$  must lie outside households' budget constraints under the equilibrium price function  $\{t(q, a), R\}$  :

$$c'_1(a, y) + t(q'(a, y), a) \geq y + \frac{y_2(q'(a, y), a) - c'_2(a, y)}{R} - \mathbb{1}_{\{q'(a, y) > 0\}} \omega \quad (15)$$

with strict inequality for a positive measure of households. Summing up across all types of household, we have

$$\begin{aligned} & \sum_a \int_y c'_1(a, y) dF_a(y) + \sum_a \int_y t(q'(a, y), a) dF_a(y) \\ & > \sum_a \int_y y dF_a(y) + \frac{\sum_a \int_y y_2(q'(a, y), a) dF_a(y) - \sum_a \int_y c'_2(a, y) dF_a(y)}{R} - (1 - \chi'(0)) \omega \\ & \geq \sum_a \int_y y dF_a(y) - (1 - \chi'(0)) \omega \end{aligned} \quad (16)$$

where the second equality follows from the second period resource constraint.

On the other hand the college's decision rule must also be inferior under the competitive price function

$$\sum_a \eta'_a(q) t(q, a) - e'(q) - \phi \leq 0$$

Summing up across all the colleges, we have that under the alternative allocation, aggregate tuition revenue cannot exceed aggregate college expenditure:

$$\sum_a \int_y t(q'(a, y), a) dF_a(y) \leq \int e'(q) d\chi'(q) + (1 - \chi'(0)) \phi$$



Thus we can show that the aggregate first period resource constraint is violated:

$$\begin{aligned}
& \sum_a \int_y c'_1(a, y) dF_a(y) + \int e'(q) d\chi'(q) + (1 - \chi'(0)) (\omega + \phi) \\
& \geq \sum_a \int_y c'_1(a, y) dF_a(y) + \sum_a \int_y t(q'(a, y), a) dF_a(y) + (1 - \chi'(0)) \omega \\
& > \sum_a \int_y y dF_a(y) - (1 - \chi'(0)) \omega + (1 - \chi'(0)) \omega \\
& = \sum_a \int_y y dF_a(y)
\end{aligned}$$

Thus, the alternative allocation is infeasible. This contradicts the original assertion. Thus, the competitive equilibrium is efficient.

### Proof of Proposition 7

First, observe that we can combine the first- and second-period budget constraints by substituting out the credit variable  $b$  to obtain the following life-time household budget constraint:

$$c_1 + \frac{c_2}{R} \leq y + \frac{y_2(q, a)}{R} - t(q, a) - \mathbb{1}_{\{q>0\}}\omega$$

Households face a college enrollment choice and an income-smoothing problem, but these two problems can be cleanly separated. Specifically, we can define

$$Y_2(a) = \max_q \left\{ \frac{y_2(q, a)}{R} - t(q, a) - \mathbb{1}_{\{q>0\}}\omega \right\} \quad (17)$$

as the maximum net income the household can generate by going to college. Then the income smoothing problem simplifies to:

$$\max_{c_1, c_2} \{ \log(c_1) + \beta \log(c_2) \}$$

s.t.

$$c_1 + \frac{c_2}{R} \leq y + Y_2(a)$$

Since equation 17 does not involve first period income  $y$ , the optimal college enrollment choice  $q(y, a)$  is independent of  $y$ .<sup>56</sup> Hence the investment model with a frictionless credit market features *no sorting by income*. The intuition is that the returns to saving in the credit market or investing in education must always be equated, irrespective of family income.

Next we explore how the college enrollment choice varies with ability. With complementarity between education and child ability in the production function (eq. 10), any efficient allocation must feature positive sorting by ability in enrollment. Formally, consider two different levels of ability, with  $a_1 > a_2$ . Suppose a child with  $a_1$  does not enroll, but one with  $a_2$  does. Now flip the enrollment pattern, letting the  $a_1$  child take the college spot of the  $a_2$  child. This is a profitable perturbation from the social planner's perspective, since the ability-education complementarity means it yields higher aggregate output. Since any competitive equilibrium is Pareto efficient (Proposition 6), this concludes the proof of Proposition 7.

Lastly, we note that the proof of the welfare theorem (Proposition 6) fails when there is a positive credit wedge  $\rho > 0$ . In that case, borrowers and lenders face different effective interest rates. Therefore, in their life-time budget constraints (eq. 15), different

---

<sup>56</sup>Note that we have implicitly assumed that tuition does not depend on income. This property can be proved using a similar logic as in Proposition 1, since the college side of the model is unchanged.

households  $(y, a)$  use different interest rates  $R(y, a)$  to discount future consumption:

$$c'_1(a, y) + t(q'(a, y), a) \geq y + \frac{y_2(q'(a, y), a) - c'_2(a, y)}{R(y, a)} - \mathbb{1}_{\{q'(a, y) > 0\}} \omega$$

This means that the derivation in eq. 16 fails, as one can no longer use the second-period resource constraint to deduce the second inequality.

### Proof of Proposition 8

In this proof we will construct a  $\bar{\rho}$  such that if the credit market friction  $\rho > \bar{\rho}$ , then the equilibrium of the investment model corresponds to that of the consumption model.

To start, we know by Proposition 3 that a competitive equilibrium exists when education is a consumption good. In that equilibrium, given the tuition function  $t(q, a)$ , a  $(y, a)$ -type household chooses his optimal allocations  $q(y, a)$  and  $c(y, a)$ . For this particular type household, we can derive an interest rate  $\hat{R}(y, a)$  at which he is unwilling to either borrow or save. We divide it into two cases:

Case 1: he enrolls in college  $q(y, a) > 0$ . In this case, the interest rate  $\hat{R}(y, a)$  is implicitly defined by the first order condition:

$$\frac{1}{y - \omega - t(q(y, a), a)} = \frac{\beta \hat{R}(y, a)}{y_2(q(y, a), a)}$$

Or

$$\hat{R}(y, a) = \frac{y_2(q(y, a), a)}{\beta [y - \omega - t(q(y, a), a)]}$$

Case 2: he does not go to college  $q(y, a) = 0$ . In this condition a similar condition implies that

$$\hat{R}(y, a) = \frac{y_2(0, a)}{\beta y}$$

Given the constructed function  $\hat{R}(y, a)$ , we can define  $\bar{\rho}$  as

$$\bar{\rho} = \max_{y, a} \hat{R}(y, a) - \min_{y, a} \hat{R}(y, a)$$

Now we need to verify two things: first  $\bar{\rho} < \infty$ ; second, for any  $\rho > \bar{\rho}$ , there exists an equilibrium with no borrowing and lending, and thus the education-as-investment model is isomorphic to the education-as-consumption model.

To see  $\bar{\rho}$  is finite, we need to show that the maximum and the minimum of  $\hat{R}(y, a)$  is finite. Observe that  $(y, a)$  lies in a closed and bounded set, and we also know that the tuition function  $t(\cdot)$  is bounded (from Proposition 3) and that the second-period income function  $y_2(\cdot)$  is bounded as well (due to the assumed functional form and because the quality support is bounded). These imply that any element  $\hat{R}(y, a)$  must be finite and therefore the maximum and the minimum must be finite (note that the extrema are always attainable as the domain is a closed set). As a result  $\bar{\rho} < \infty$ .

To see the second property, fix any  $\rho > \bar{\rho}$ . Now suppose, to the contrary, that in equilibrium there is strictly positive borrowing and lending at some equilibrium interest rate

R. Let  $(y_1, a_1)$  be the household with positive lending. Then it must be the case that  $R > \hat{R}(y_1, a_1)$ , thus the borrowing rate in this economy must be

$$\begin{aligned}
 R + \rho &> \hat{R}(y_1, a_1) + \rho \\
 &> \hat{R}(y_1, a_1) + \bar{\rho} \\
 &= \hat{R}(y_1, a_1) + \max_{y,a} \hat{R}(y, a) - \min_{y,a} \hat{R}(y, a) \\
 &\geq \hat{R}(y_1, a_1) + \max_{y,a} \hat{R}(y, a) - \hat{R}(y_1, a_1) \\
 &= \max_{y,a} \hat{R}(y, a)
 \end{aligned}$$

As the borrowing rate exceeds the maximum interest rate at which *any* households would like to borrow, there will be no demand for borrowing in this economy. A contradiction. Thus, there will be no credit in this equilibrium.

## 8.2 Computation Appendix

### 8.2.1 Algorithm for Baseline Calibration

This section explains the computational algorithm used to solve the quantitative model with two ability types. The key equilibrium object to solve for is the college distribution function  $\chi(\cdot)$  defined over a discrete grid on college quality  $q$ . Note that this equilibrium exists by step 1 in the proof of Proposition 3. Hence we are computing an approximation of an exact equilibrium, rather than an approximate equilibrium. The algorithm uses extensively the sorting property of the model, i.e., richer households are always more likely to go to college and, in case they do, always prefer better college quality. However, a complication arises regarding the need-based aid: households with income level just above the threshold  $y^*$  might have less incentive to attend college than households with income level just below. The way we deal with it is to assume that households with and without need-based aid are two different types. Then within each type, sorting by income holds. To highlight the working of the algorithm, in the following we lay out an algorithm with just two types (high and low ability). Details on computation with more types including different residence status and pell eligibility are available on request.

1. Construct a grid on college quality  $q$  with values  $q(1), q(2), \dots, q(N)$  where  $q(1) > 0$  and  $q(N) = q_{max}$ .
2. Make an initial guess of the share of high-ability students not entering college:  $\eta(0)$ . By definition, the fraction of low-ability students not entering college is  $1 - \eta(0)$ .
3. Solve for  $\chi(0)$  from the zero profit condition of colleges of quality  $q(1)$ .
  - (a) Starting with a conjecture for  $\chi(0)$ , compute the income of the “marginal” household attending colleges of quality  $q(1)$ :

$$\begin{aligned}y(i^h(1)) &= y(\eta(0)\chi(0)) \\y(i^l(1)) &= y((1 - \eta(0))\chi(0)).\end{aligned}$$

Next, pin down the college tuition  $(t^h(1), t^l(1))$  of the  $q(1)$  college by the marginal household's indifference condition:

$$\begin{aligned}\log(y(i^h(1))) + \varphi \log(\kappa) &= \log(y(i^h(1)) - t^h(1)) + \varphi \log(\kappa + q(1)) \\ \log(y(i^l(1))) + \varphi \log(\kappa) &= \log(y(i^l(1)) - t^l(1)) + \varphi \log(\kappa + q(1)).\end{aligned}$$

Given the prevailing market tuition  $t^h(1)$  and  $t^l(1)$ , solve the  $q(1)$  college optimization problem and obtain its profit  $\pi(1)$  as well as its optimal decision rules  $\eta(1), e(1)$ .

(b) Use the mapping described in part (a) to solve for  $\chi(0)$  such that  $\pi(1) = 0$ .

i. Check the value of  $\pi(1)$  at boundaries  $\left[ \chi_{lb}(0) = 0; \chi_{ub}(0) = \min \left\{ \frac{1}{\eta(0)}, \frac{1}{1-\eta(0)} \right\} \right]$ .

The upper bound arises because the total mass of high(low) ability is 1.

Note that the profit  $\pi(1)$  should be increasing in  $\chi(0)$ , as the market tuition rates  $t^h(1)$  and  $t^l(1)$  are both increasing in  $\chi(0)$ .

A. If  $\pi(1) > 0$  at  $\chi_{lb}(0) = 0$ , or  $\pi(1) < 0$  at  $\chi_{ub}(0) = \min \left\{ \frac{1}{\eta(0)}, \frac{1}{1-\eta(0)} \right\}$ , zero profits cannot be obtained at grid  $q(1)$ . Thus we delete  $q(1)$  and set  $q(1) = q(2)$  and go back to step 3, else go to step ii.

ii. As  $\pi(1) < 0$  at  $\chi_{lb}(0) = 0$  and  $\pi(1) > 0$  at  $\chi_{ub}(0)$ , one can solve for  $\chi(0)$  from  $\pi(1) = 0$  using a simple one-dimensional nonlinear solver.

4. Having solved for  $\{\chi(i)\}_{i=1}^{n-1}$ , along with  $\{\eta(i), e(i)\}_{i=1}^n$  we now solve for  $\chi(n)$  from  $\pi(n+1) = 0$ .

(a) Starting from a conjecture for  $\chi(n)$ , compute the income of the marginal household attending colleges of quality  $q(n+1)$ :

$$\begin{aligned} y(i^h(n+1)) &= y\left(\sum_{i=0}^n \eta(i) \chi(i)\right) \\ y(i^l(n+1)) &= y\left(\sum_{i=0}^n (1-\eta(i)) \chi(i)\right). \end{aligned}$$

Next, pin down the college tuition  $(t^h(n+1), t^l(n+1))$  of the  $q(n+1)$  college from the following household first-order conditions:

$$\begin{aligned} t^h(n+1) &= \left[ 1 - \left( \frac{\kappa + q(n)}{\kappa + q(n+1)} \right)^\varphi \right] y(i^h(n+1)) + \left( \frac{\kappa + q(n)}{\kappa + q(n+1)} \right)^\varphi t^h(n) \\ t^l(n+1) &= \left[ 1 - \left( \frac{\kappa + q(n)}{\kappa + q(n+1)} \right)^\varphi \right] y(i^l(n+1)) + \left( \frac{\kappa + q(n)}{\kappa + q(n+1)} \right)^\varphi t^l(n). \end{aligned}$$

Given the prevailing market tuition  $t^h(n+1)$  and  $t^l(n+1)$ , solve the  $q(n+1)$  college optimization problem (procedure outlined below) and obtain its profit  $\pi(n+1)$  as well as the optimal decision rules  $\eta(n+1), e(n+1)$ .

(b) Given the mapping described in part (a), solve for  $\chi(n)$  such that  $\pi(n+1) = 0$ .

- i. Check the value of  $\pi(n)$  at boundaries  $\left[ \chi_{lb}(n) = 0; \chi_{ub}(n) = \min\left(\frac{1-i^h(n)}{\eta(n)}, \frac{1-i^l(n)}{(1-\eta(n))}\right) \right]$ .
  - A. If  $\pi(n+1) > 0$  at  $\chi_{lb}(n)$ , this implies that  $q(n+1)$  college would always make strictly positive profits and keep growing, squeezing  $q(n)$  college out of the market. Thus, we can delete the grid point  $q(n)$ . Set  $q(n) = q(n+1)$  and go back to step 3 with the new grid on  $q$ .
  - B. If  $\pi(n+1) < 0$  at  $\chi_{ub}(n)$ , this implies that  $q(n+1)$  college always makes negative profits and thus is driven out of the market. Thus, we delete the grid point  $q(n+1)$ . Set  $q(n+1) = q(n+2)$  and go back to step 3 with the new grid on  $q$ .
  - C. If  $\pi(n+1) < 0$  at  $\chi_{lb}(n)$  and  $\pi(n+1) > 0$  at  $\chi_{ub}(n)$ , then we can solve for  $\chi(n)$  such that  $\pi(n+1) = 0$ .

5. Having solved for  $\{\chi(i)\}_{i=1}^{N-1}$ , along with  $\{\eta(i), e(i)\}_{i=1}^N$ , we still have  $\chi(N)$  undetermined. We pin it down using the consistency requirement for high ability spots at  $q(N)$  colleges:

$$\chi(N)\eta(N) = 1 - i^h(N).$$

Lastly, check the consistency requirement for low ability spots at  $q(N)$  college:

$$1 - i^l(N) - \chi(N)(1 - \eta(N)) = 0.$$

If this requirement is satisfied (to desired numerical accuracy), stop. If not, go back to step 2 and adjust  $\eta(0)$ .

## 8.2.2 More Ability Types: Computation

We first present a computational algorithm that can be used to solve a model with more than two ability types.<sup>57</sup> We then apply the algorithm to solve a model with 10 ability types and compare the resulting equilibrium college distribution with our baseline calibration with 2 ability types. We find that varying the number of grid points has negligible effects on the equilibrium quality distribution and on key moments of the enrollment and tuition distributions.

---

<sup>57</sup>More details are available upon request.



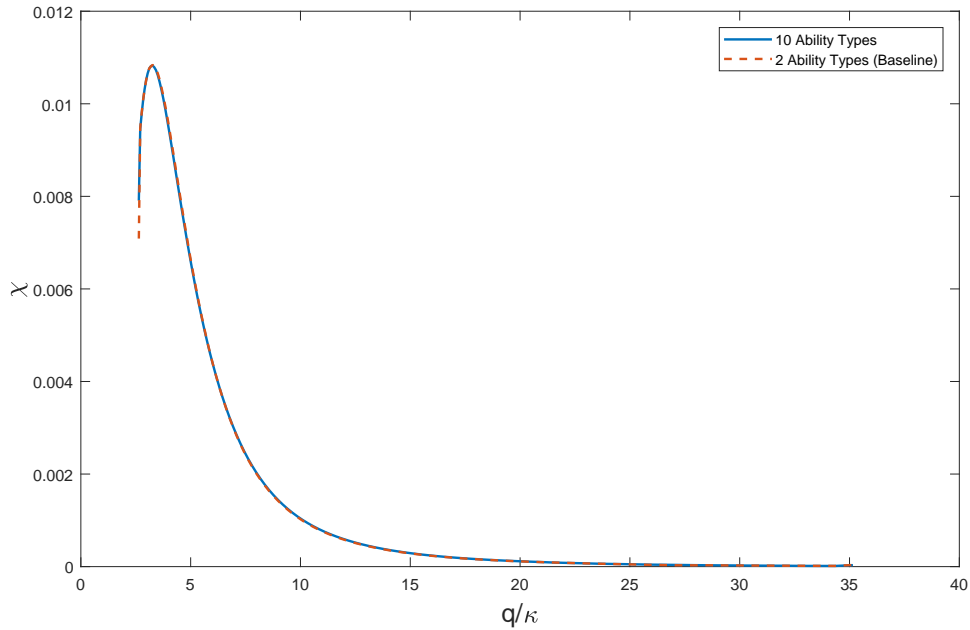
With the college problem reformulated as in eq. 7, we can simplify the college market-clearing condition, replacing eq. 4 with the following two conditions:

$$\begin{aligned}\chi(Q) &= \sum_r \mu_r \sum_a \mu_a \int \mathbb{1}_{\{q(y,a,r) \in Q\}} dF_a(y) \quad \forall Q \subset \Omega, \\ \int_Q \bar{a}(q) d\chi(q) &= \sum_r \mu_r \sum_a \mu_a \int \mathbb{1}_{\{q(y,a,r) \in Q\}} a dF_a(y) \quad \forall Q \subset \Omega.\end{aligned}$$

The first condition states that the measure of students in any quality set  $Q$  is consistent with student attendance choices. The second equates the average student ability demanded by colleges producing in quality set  $Q$  to the average ability of the students choosing to supply to quality set  $Q$ . This greatly simplifies the set of market clearing conditions that needs to be checked. The general strategy is to solve for a college distribution over a quality grid where we know with certainty that colleges are active, and to then check for profitable entry at the bottom.

1. Set up a grid of college quality  $(q_1, q_2, \dots, q_N)$  where we know with certainty that colleges enter.
2. Make an initial guess of the vector of corresponding discount rates  $D_0 = (d_1, d_2, \dots, d_N)$ .
3. Given the discount rates, use the college first-order conditions and zero profit conditions to compute the set of implied baseline tuition  $(b_1, b_2, \dots, b_N)$ . Thus, we obtain a full set of tuition schedules.
4. Given the tuition schedules, use the household's indifference condition to pin down a set of income thresholds for each ability type  $(y_1^a, y_2^a, \dots, y_N^a)$ , where  $y_i^a$  is the income of a household indifferent between quality  $i - 1$  and quality  $i$  colleges.
5. Given the income thresholds, compute the supply of average ability to each college  $(a_1^s, a_2^s, \dots, a_N^s)$ .
6. Given demand for ability by each college  $i$  (pinned down given  $d_i$  from the college first-order conditions)  $(a_1^d, a_2^d, \dots, a_N^d)$  check market clearing:  $a_n^d - a_n^s = 0 \forall n$ .
7. If markets do not clear, go back to step 2 and adjust the discounts  $d_n$ .
8. Check for profitable entry at the bottom. For instance, suppose a college of quality  $q_0 < q_1$  enters. To charge the maximum tuition, it has to appeal to the marginal household with income  $y_1^a$ . Thus, we can use the household's indifference condition

Figure 11: College Distribution



to pin down tuition  $t_0^a$ . We then solve the college problem and check its profit. If profit is positive, go back to step 1 and add  $q_0$  to the grid of college quality. Otherwise, we stop.

We now solve the college model with 10 points in the grid on ability and compare the results to our benchmark two-ability-types calibration. We use the same parameterization as in our baseline. We discretize the 10 grid points such that 1) the 10-grid-point model has the same variance of ability as the baseline, and 2) the conditional mean of income distribution varies linearly with ability with the same slope as the baseline. Figure 11 plots the equilibrium college distribution (density) with different number of ability types.

### 8.3 Data Appendix

In this data appendix, we first explain how we construct first moments for the 2016 and 1990 calibrations. We then explain how we construct the second moments in Table 2 from micro college-level data.

#### Enrollment and graduation rates

We target observed 4 year college enrollment rates and impose observed dropout rates by ability. The graduation rate from the bottom half of the AFQT distribution is 0.52 (Hendricks et al., 2018) while the graduation rate from the top half of the AFQT distribution is 0.78. From the NLSY97, we computed that 12.5% of below median ability children have graduated with a college degree. That suggests  $0.125/0.52 = 24.0\%$  of low ability children enrolled in 4 year colleges. Similarly, 52.8% of above median ability children have a college degree, suggesting  $0.528/0.78 = 67.7\%$  enrolled. The total enrollment rate for the economy is therefore  $0.5 \times 0.240 + 0.5 \times 0.677 = 45.85\%$ .

However, note that the NLSY calculation suggests an overall graduation rate of  $(0.125 + 0.528)/2 = 32.65\%$  which is lower than the 36.1% number from the CPS for 25-29 year-olds. We target the CPS number, because we can measure that on a consistent basis over time. So to match the CPS number we adjust the NLSY-based graduation rates by a factor of  $36.1/32.65 = 1.106$ . So that implies  $1.106 \times 52.8 = 58.4\%$  of high ability children graduate, and  $1.106 \times 12.5 = 13.8\%$  of low ability children graduate, and the corresponding enrollment rates are  $0.584/0.78 = 74.9\%$  and  $0.138/0.52 = 26.5\%$ . Thus, the overall enrollment rate we target is  $(74.8\% + 26.5\%)/2 = 50.7\%$ .

#### Shares in public vs private colleges, and shares in-state vs out-of-state

In 2016, 78% of public 4 year college students were from in-state (Trends in College Pricing (TICP) 2018, Fig 23). Enrollment in 4 year schools in public and private colleges in 2016 was, respectively,  $4,994,668 + 1,185,002$ , and  $2,187,122 + 466,900$ , giving a public share of

$$\frac{4994668 + 1185002}{4994668 + 1185002 + 2187122 + 466900} = 70.0\%.$$

(Trends in College Pricing 2018, Fig 21). So the share of in-state students as a share of all students in 4 year colleges is  $0.78 \times 0.70 = 54.6\%$ .

#### Ability discount (institutional aid)

Average institutional aid at public schools in 2015-16 was \$2,274 (Trends in College Pricing 2018, Fig 18) and \$14,055 at private schools (Trends in Student Aid (TISA) 2018, Fig 19). So average per student institutional aid is  $0.7 \times 2,274 + (1 - 0.7) \times 14,055 = \$5,808$ .

### Need based aid

In 2016-17, 32% of students were receiving a Pell grant (Trends in Student Aid 2018, Fig 20A), and the average grant was \$3,800 (Trends in Student Aid 2018, Fig 21A). So the unconditional average was  $0.32 \times 3,800 = \$1,216$ .

In addition, there are need-based state grants. Of all state grant aid, 76% is need based (Trends in Student Aid 2018, Fig 23A), and state grant aid is 1,442 at public 4-year colleges (TISA 2018 Fig 18 for 15-16) and 944 at private 4-year colleges (TISA 2018 Fig 19 for 15-16). So a good target number for need-based state grant aid is  $0.76 \times (0.7 \times 1,442 + (1 - 0.7) \times 944) = 982$ . So total need-based aid per student (Pell grants + state need-based aid) is  $1,216 + 982 = \$2,198$ .

### General subsidy to students $p_0$

The general subsidy is average sticker tuition minus average net tuition minus institutional aid minus public need-based aid =  $\$19,152 - \$9,249 - \$5,808 - \$2,198 = \$1,896$ .

### In-state vs. out-of-state tuition

Sticker tuition numbers for 2016-2017 are \$9,650 for public in-state and \$24,930 for public out-of-state (TICP 2016, Table 1A). So in-state tuition is 38.7% of out-of-state tuition. The in-state subsidy per subsidized student is \$15,280, which translates to \$8,343 per student overall.

### Fixed costs net of subsidies directly to colleges

We have dealt separately with all subsidies that go to students directly. There are separate subsidies that go to colleges and benefit students indirectly. In the model, a fixed subsidy per student that goes to schools is the same (in terms of allocations and welfare) as a subsidy that goes to students. But whereas net tuition is the same in both cases, sticker tuition will be different: a subsidy that goes to schools will reduce sticker tuition, while a subsidy that goes to students will not.

We know that in aggregate the following aggregate budget constraint for the university sector must be satisfied

$$E + \phi = \text{Tuition Revenue} + \text{Other Revenue}$$

where  $\phi$  is the fixed administration cost per student, and  $E$  is per student variable spending, which we define as instructional spending + student services spending.

College revenue from tuition is

$$\text{Tuition Rev.} = \text{Sticker Tuition} - \text{Ability Discounts (Inst. Aid)} = \text{Net Tuition} + \text{Public Student Aid}$$

where here sticker tuition is already net of the in-state discount for in-state students, and public student aid  $p(y)$  is general plus need-based public student aid.

Other revenue is

$$\text{Other Revenue} = \bar{s} + \text{In-state Transfers}$$

where in-state transfers are the transfers from the state for taking in-state students, and  $\bar{s}$  are other sources of subsidies that go straight to colleges (rather than to students). Thus,

$$E + \phi = \bar{s} + \text{In-state Transfers} + \text{Net Tuition} + \text{Public Student Aid.}$$

Now in the data we can measure  $E$ , In-state Transfer, Net Tuition, and Public Student Aid. From these we can get

$$\phi - \bar{s} = \text{In-state transfers} + \text{Net Tuition} + \text{Public student aid} - E$$

In particular, in 2016,

- Expenditure  $E = 0.7 \times (12,539 + 2107) + 0.3 \times (17,996 + 4,753) = 17,077$  (NCES 334.10 and 334.30).
- Public student aid  $= 1,896 + 2,198 = 4,094$ .
- In-state transfers  $= 0.78 \times 0.7 \times (24,930 - 9,650) = 8,343$ .
- Net Tuition  $= 9,249$ .

Hence:

$$\begin{aligned} \phi - \bar{s} &= 8,343 + 9,249 + 4,094 - 17,077 \\ &= 4,610 \end{aligned}$$

Thus, fixed costs net of general subsidies to colleges are positive.

#### Forgone Earnings

From the CPS (Series ID: LEU0252886300) we take median usual weekly earnings for full-time wage and salary workers aged 16-24. In current dollars, this was \$259 in 1990 and \$501 in 2016. Adjusted by the CPI and assuming 20 weeks of college per year, forgone earnings from attending college is \$10,020 in 2016 dollars in both 2016 and 1990.

We now describe how we construct all the data moments reported in Table 3 for 1990.

#### Enrollment and Tuition in 1990

The share of 25-29 year-olds with a college degree was 23.3% in the CPS in 1990.<sup>58</sup> Assuming the same graduation to enrollment ratio as 2016, the implied enrollment rate for 1990 is 32.7%.

Of the students enrolled, we assume 70% were in public colleges, as in 2016 (according to NCES Table 303.70, the 1990 public share was 70.7%). Of the students in public colleges, we assume 83% were in-state students in 1990 (the earliest estimate for this number we found was 83% for 2004, TICP 2016, Fig 22). This implies a share of in-state students of  $0.7 \times 0.83 = 0.581$ .

We estimate average net tuition in 1990 to be \$6,034. Average net tuition at private colleges was \$11,750, and average net tuition for in-state students at public colleges was \$2,000. Our estimate for net tuition for out-of-state students at public colleges is sticker tuition for this group (\$12,837, see below) minus the gap between sticker and net tuition for in-state students (i.e.,  $\$3,520 - \$2,000 = \$1,520$ ) which implies a net tuition value of \$11,317. Overall average net tuition is a weighted average of these three groups of students:

$$\text{Net Tuition} = 0.7 * 0.83 * 2,000 + 0.7 * (1 - 0.83) * 11,317 + 0.3 * 11,750 = \$6,034$$

#### Subsidies in 1990

Need-based aid: The average Pell grant for 1990-1991 was \$2,720 per recipient (Trends in Student Aid 2018, Fig 21A). We use the 1990-91 Federal Pell Grant Program End of Year Report to estimate the share of students receiving Pell grants in 1990. Applications were 63% of students, of which 63.1% were found eligible, of which 75.5% actually ended up receiving money, suggesting 30% of students were getting grants. Equivalently 3.4 million out of 11.4 million students = 30%. Thus the unconditional Pell grant average for 1990-91 is  $0.30 \times 2,720 = \$816$ . We next construct a state need-based grant estimate for 1990-1991. From 1990-91 total state grant aid per student has grown by 105%, and need-based aid has grown by 75.2% (TISA 2018, Fig 23A). Thus, our state need-based grant estimate for 1990-91 is  $982/1.752 = 561$  per student. Total need-based aid in 1990 is thus  $816 + 561 = \$1,377$ .

General subsidies to students,  $p_0$ : We estimate this as 12% of state grant aid in 1990 (88% of such aid is need-based) which translates to \$76.

Total public aid to students: Need-based aid plus the lump-sum component  $p_0$ . Thus, total public student aid was  $\$1,377 + \$76 = \$1,453$ .

In-state transfers: We estimate sticker tuition for out-of-state students at public col-

---

<sup>58</sup>Table A-2 is available here: <https://www.census.gov/data/tables/time-series/demo/educational-attainment/cps-historical-time-series.html>

leges in 1990 (for which we could not find data) as sticker tuition for private college students in 1990 (\$17,240) times the ratio of sticker tuition for out-of-state public relative to private students in 2016, which is \$24,930 divided by \$33,480, which translates to an estimate for out-of-state sticker tuition in 1990 of \$12,837. Thus, in-state transfers, per in-state student are  $\$12,837 - \$3,520 = \$9,317$ , where \$3,520 is public in-state sticker tuition in 1990. Thus, on a per-student basis, the value of transfers to support in-state students was  $\text{In-state Transfers} = 0.83 \times 0.7 \times \$9,317 = \$5,413$ .

General subsidies to colleges  $\bar{s}$ : To estimate general per student subsidies to colleges in 1990, we use the aggregate budget constraint for the college sector, as for 2016. To do so we require an estimate for variable expenditure per student in 1990.

We measure variable expenditure (defined as instructional expenses+student services) in 1990 in the same way as in 2016. The difficulty is that the NCES changed its reporting standards twice during late 1990s and early 2000s, making numbers not directly comparable across 1990 and 2016. We instead compute growth rates in each subperiod during which reporting standards remained consistent, and use these growth rates to infer 1990 expenditure. For public colleges, the cumulative growth of variable expenditure was 17.9 percent from 1990 to 2001. The growth rate was 16 percent from 2003 to 2014. Given that the variable expenditure for public colleges was \$11,881 in 2014-2015, we infer that consistently defined variable expenditure in 1990 was \$8,680. For private nonprofit colleges, the growth rate was 11.3 percent between 1990 and 1996. The growth rate was 10.4 percent between 1996 and 1999. The growth rate was 7.4 percent between 1999 and 2003. The growth rate between 2003 and 2014 was 13.5 percent. Given these growth rates and variable expenditure for private colleges of \$22,120 in 2014-2015, the estimated variable expenditure for private nonprofit colleges in 1990 is \$14,757. Putting all this together, we estimate per student expenditure in 1990 of  $E = 0.7 \times 8,680 + 0.3 \times 14,757 = 10,503$ .

Hence, 1990 fixed costs net of general subsidies to students are

$$\begin{aligned} \phi - \bar{s} &= 5,413 + 6,034 + 1,453 - 10,503 \\ &= 2,396 \end{aligned}$$

### 8.3.1 Statistics of Table 2

The second moments are computed from College Scorecard microdata merged with the Mobility Report Cards data set (Chetty et al. 2020), which has higher-quality household income data. Specifically, we download the most recent data from the College Scorecard (<https://collegescorecard.ed.gov/data/>) and merge it with household income data from

the Mobility Report Cards online data Table 2.<sup>59</sup> Next, we describe how we construct each variable.

Sticker tuition: In-state tuition and fees (variable name `tuitionfee_in`).

Net tuition: Average net price paid (`NPT4_pub` for public colleges and `NPT4_priv` for private colleges). Note this measure includes the full cost of attendance (including living expenses). We construct a measure of living expenses by subtracting tuition and fees (`tuitionfee_in`) from the full cost of attendance (`costt4_a`). Net tuition is then obtained by subtracting the living expenses from the average net price paid.

Household income: From Mobility Report Cards online data Table 2 (variable name `par_mean`). Mean income is \$87,335.

Fraction of high ability: We collect data on national averages of the SAT score (when- ever the SAT score is not available, we substitute the ACT score.) We then assume that the score is normally distributed at each college and use the college-specific 25th percentile and 75th percentile SAT score (`satmt25`, `satvr25`, `satwr25`, `satmt75`, `satvr75`, `satwr75`) to back out the mean and variance of the distribution at each college. Then we compute the fraction of high-ability students at each college as the fraction with a score higher than the national average.

---

<sup>59</sup>Note that to link online data Table 2 to College Scorecard data, we need to use online data Table 11.